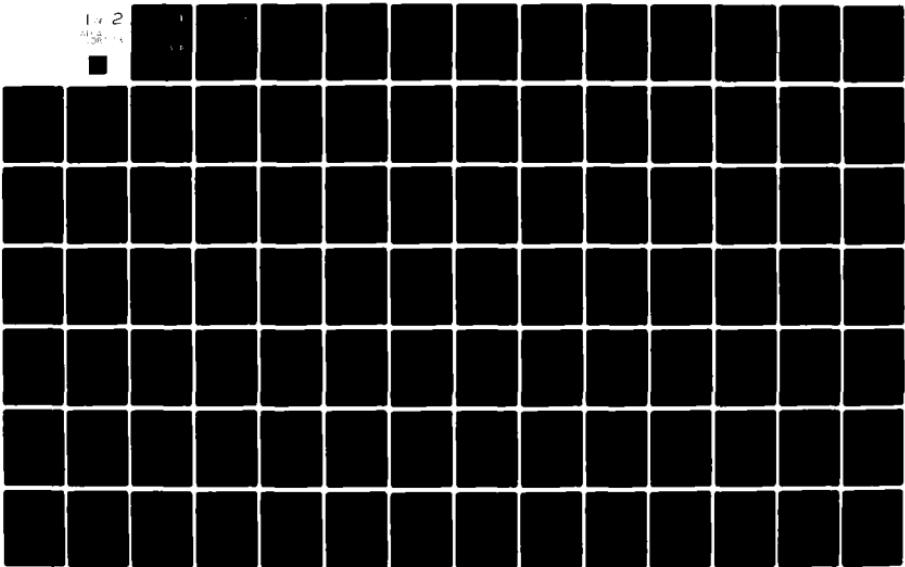


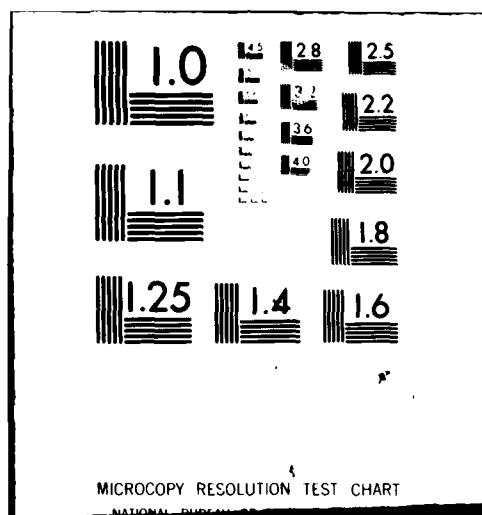
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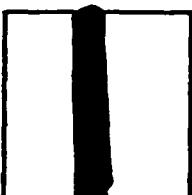
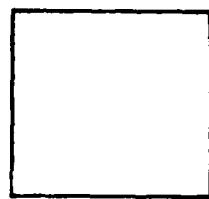




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# DAVID W. TAYLOR NAVAL SHIP RESEARCH AND DEVELOPMENT CENTER

Bethesda, Md. 20084



MODIFICATIONS TO COMPUTER  
PROGRAM FOR PARAMETER ESTIMATION  
FOR THE GENERALIZED GAMMA DISTRIBUTION

by

Michel K. Ochi

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FOR THE GENERALIZED GAMMA DISTRIBUTION

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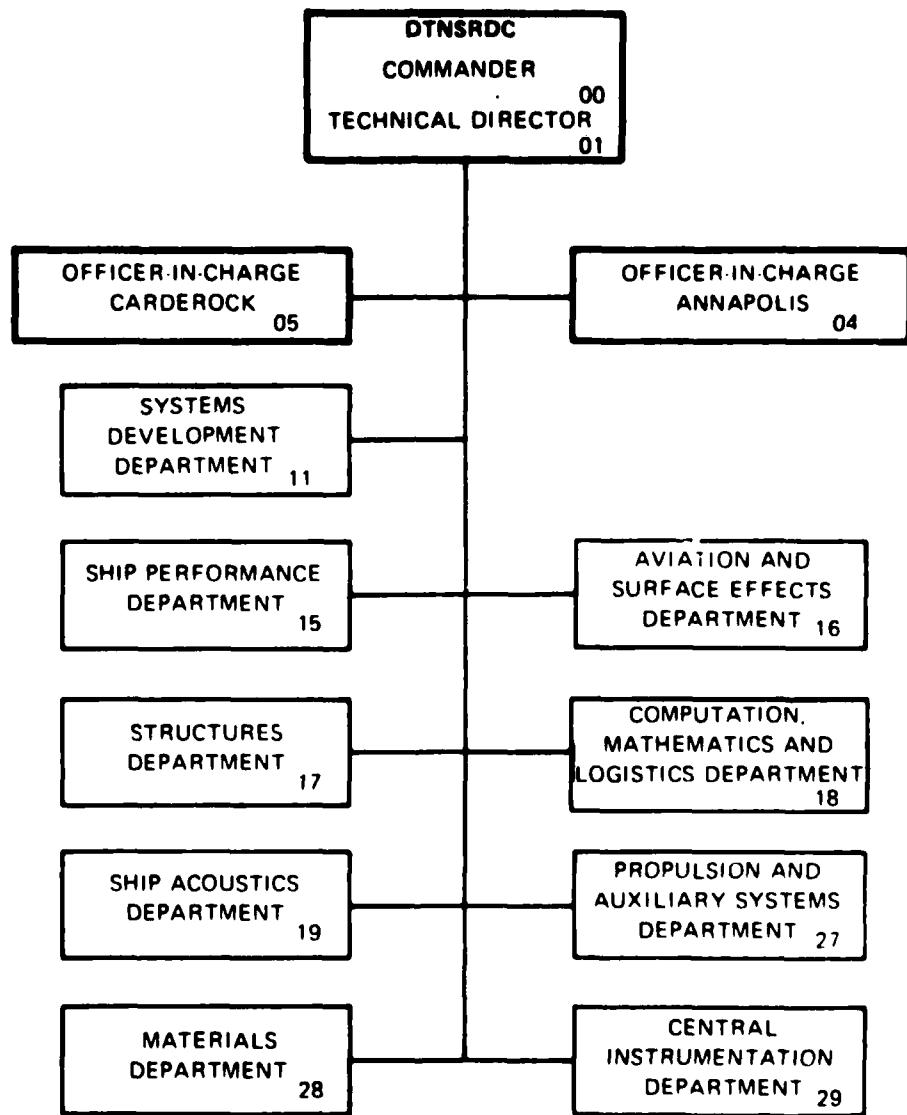
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the Stacy-Mihram method are investigated. The most successful alternative method is the maximum likelihood method, which is incorporated into the new digital computer program. A program listing and an example of the generated output are included in this report.

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## ABSTRACT

An extensive study of an existing digital computer program for determining parameters required to represent SES 100B trials data by a generalized gamma function is carried out. The purpose of the study is to resolve the difficulties encountered in earlier studies. In these earlier studies, approximately 5 percent of the cases analyzed produced unrealistic values for the distribution parameters. In order to resolve this problem, data sensitivity studies are carried out. Based on results of the study, criteria are developed for determining cases that should not be processed. Three other techniques besides the Stacy-Mihram method are investigated. The most successful alternative method is the maximum likelihood method, which is incorporated into the new digital computer program. A program listing and an example of the generated output are included in this report.

## ADMINISTRATIVE INFORMATION

This work was conducted at the David W. Taylor Naval Ship Research and Development Center and authorized by Naval Sea Systems Command, Surface Effect Ships Project Office, Program Element Number 63534N. It is identified as Work Unit Number 1-1506-012.

## INTRODUCTION

A computerized procedure for evaluating the parameters of the generalized gamma distribution from a set of random data following the method proposed by Stacy and Mihram (see Reference 1) has been previously

developed and extensively exercised (see Reference 2 and 3). During these earlier studies it was discovered that what appeared to be unrealistic values for the distribution parameters were obtained in roughly 5 percent of the cases analyzed. The initial purpose of the study, whose results are summarized in this report, was to explore in detail a representative sample of cases resulting in unrealistic values of the parameters, and if possible, to develop an alternate method for processing these cases.

The results of the present investigation can be summarized as follows:

- a. In the many cases studied, it appears that the predicted extreme design values are most sensitive to experimental measurements and not the numerical procedure for determining the parameters. This is particularly true when one of the parameters is negative.
- b. All four methods considered in this investigation yield probability distribution functions which represent very well the experimentally obtained histogram even though some of the parameter values were judged unrealistic in previous studies. However, the Stacy-Mihram and maximum likelihood methods are by far the most efficient and are the preferred procedures based on probabilistic arguments.
- c. In some cases the predicted extreme design values are unrealistically large whereas the significant values are

realistic. Based on an analysis of over 100 cases, it can be concluded that cases having a nondimensional third logarithmic moment greater than -0.5 produce unrealistic extreme design values. Results of a data sensitivity study have revealed that in all these cases the data may not be considered as a sample taken from a steady-state random phenomenon, and hence it is recommended not to carry out statistical analysis for these cases.

## MATHEMATICAL FORMULATION

### OVERVIEW

The generalized gamma density function under consideration in this report has the form

$$f(x; \lambda, m, c) = \frac{c \lambda^m}{\Gamma(m)} x^{m-1} e^{(-\lambda x)^c} \quad (1)$$

where  $\lambda$ ,  $m$ , and  $c$  are estimation parameters and  $\Gamma(m)$  is the gamma function. Many other distributions such as the exponential, gamma, chi-squared, Weibul, hydrograph, chi, truncated normal, and Rayleigh are special cases of the generalized gamma function. Due to the general nature of  $f(x; \lambda, m, c)$ , it is the preferred distribution particularly when the exact form of the distribution function is unknown.

There are several techniques that can be used to estimate the three parameters. Four such techniques are considered in this effort. These are the Stacy-Mihram, the maximum likelihood, the nonlinear least squares, and grid search methods. The last two methods determine the parameters by searching for that set of parameters resulting in a minimum error between the experimental and theoretical histograms. Since statistical inference cannot be applied in a rigorous manner, the last two methods are not recommended. The Stacy-Mihram method determines the parameters by requiring that the first three theoretical logarithmic moments agree with the logarithmic moments determined from the experimental data. The maximum likelihood method determines parameters such that the joint probability density function is a maximum.

A detailed derivation of the Stacy-Mihram method is presented in Reference 2 pages 6 - 8. A derivation of the nonlinear least squares procedure is presented in Reference 4. The grid search procedure merely involves a systematic search of the  $\lambda$ ,  $m$ ,  $c$  space for parameters that minimize the error between the predicted distribution and the measured distribution. A derivation of the equations required in the maximum likelihood parametric estimation method is presented in the next section.

#### MAXIMUM LIKELIHOOD METHOD

Briefly, the maximum likelihood estimation technique involves finding the parameter (or parameters) which maximizes the joint probability density function of a random sample  $x_1, \dots, x_N$  from a distribution with  $f(x, \theta)$  as its probability density function and  $\theta$  as an element of the parameter space. The joint probability density function, referred to as the likelihood function, is given by

$$L = \prod_{i=1}^N f(x_i; \theta) \quad (2)$$

Assume one has  $N_i$  samples at random variable  $x_i$ . Then the maximum likelihood function will have the form

$$L = \prod_i f(x_i; \theta)^{N_i} \quad (3)$$

The natural logarithm of the likelihood function is used in maximizing since it is a maximum when the function itself is a maximum, and it proves more convenient to use. The logarithm is given by,

$$\ln L = \sum_{i=1}^k N_i \ln f(x_i) \quad (4)$$

The maximum of  $\ln L$  is obtained by searching for parameters  $\theta$  such that,

$$\left. \frac{\partial \ln L}{\partial \theta_i} \right|_{i=1, N_p} = 0.0 \quad (5)$$

where  $N_p$  = the number of parameters.

For the generalized gamma function given by Equation (1), the likelihood function becomes

$$L = \prod_{i=1}^k \left[ \frac{c \lambda^m x_i^{m-1}}{\Gamma(m)} e^{(-c \lambda x_i)^m} \right]^{N_i} \quad (6)$$

and  $\ln L$  becomes

$$\begin{aligned} \ln L = & N \ln c + cm N \ln \lambda - N \ln \Gamma(m) \\ & + (cm-1) \sum_i N_i \ln x_i - \lambda^c \sum_i x_i^c N_i \end{aligned} \quad (7)$$

In the above equation,  $c$  is assumed positive. If the above equation is differentiated with respect to  $cm$ ,  $\lambda$  and  $m$  and set to zero, the following equations can be derived (see Reference 5).

$$\lambda = \left[ (m N) / \sum_i x_i^c N_i \right]^{1/c} \quad (8)$$

$$\ln c + \ln \left( \sum_i x_i^c N_i \right) - \ln c m - \ln N + \Psi(m) - c \sum_i N_i \ln x_i / N = 0 \quad (9)$$

$$1/m + \Psi(m) - \ln c m + \ln c + \ln \left( \sum_i x_i^c N_i \right) - \ln N - c \sum_i \left( \ln x_i (x_i^c N_i) \right) / \sum_i x_i^c N_i = 0 \quad (10)$$

where the summation on  $i$  is from 1 to  $K$ . An expression for  $m$  can be obtained from the above equations:

$$1/m = c \left[ \frac{\sum_i (\ln x_i) (x_i^c N_i)}{\sum_i x_i^c N_i} - \frac{\sum_i N_i \ln x_i}{N} \right] \quad (11)$$

The parameter  $c$  can be obtained by solving the following equation numerically.

$$\ln c + \ln \left( \sum_i x_i^c N_i \right) + \ln \left[ \frac{\sum_i (\ln x_i) (x_i^c N_i)}{\sum_i x_i^c N_i} - \frac{\sum_i N_i \ln x_i}{N} \right] \quad (12)$$

$$- \ln N + \Psi \left[ c \left( \frac{\sum_i (\ln x_i) (x_i^c N_i)}{\sum_i x_i^c N_i} - \frac{\sum_i N_i \ln x_i}{N} \right) \right] - \frac{c \sum_i N_i \ln x_i}{N} = 0$$

Once a value of  $c$  is determined by solving Equation (12), values for  $\lambda$  and  $m$  can be obtained via Equations (8) and (11). The numerical method used to solve Equation (12) for  $c$  is due to Wegstein (see Reference 5).

#### ANALYSIS

During previous studies in which the Stacy-Mihram method was used to process SES 100B data, large design extreme values were predicted in approximately 5 percent of the cases although the most significant values agreed with the experimental measurements. This difficulty seemed to be associated with large values of the parameter  $m$  and with negative values of the parameter  $c$ . In order to resolve this difficulty, several cases were examined in detail. The results of this investigation follows.

The first case considered in this investigation is a set of pitching motion data measured during the SES 100B full scale trials program at 30 kts in a seastate of 3. This particular case, identified as AR1 3E 30, is presented in Reference 3. A comparison of the estimated probability density function with the experimental histogram is presented in Figure 1. As is demonstrated by the results in the figure, the agreement is excellent. All other cases examined show similar results. Assuming that an objective of the estimation procedure is to develop the best representation of the data that is possible, it is concluded that large values of  $m (> 5)$  are not necessarily unreasonable.

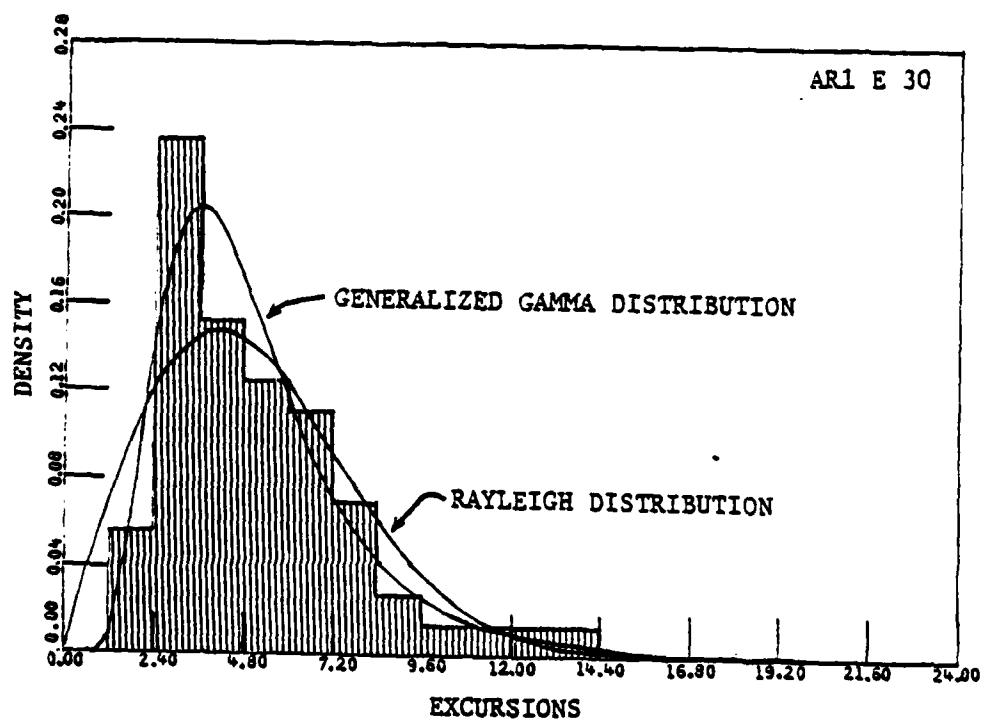


FIGURE 1. A COMPARISON OF THE GENERALIZED GAMMA DISTRIBUTION WITH EXPERIMENTAL DATA

Also, since the predicted representations always represent the data very well for both large and small  $m$ , the numerical procedure used to solve the transcendental equations for the estimation parameters is judged sufficiently accurate.

Further investigations revealed that in some of the cases the predicted design extreme values were unreasonably high as compared with the most significant values. In order to determine more precisely a reasonable value for the ratio of the design extreme value to the most significant value ( $X_D/X_S$ ), an analysis of over 100 previously processed data cases were examined. The results indicated that unrealistically large ratios were obtained for those cases for which the normalized logarithmic moment (see Equation 2.4 in Reference 2) is greater than -0.5. Most significant values were realistic as compared with those measured in trials. Consequently, in the current data base, cases for which  $T > -0.5$  should not be processed to determine design extreme values using the generalized gamma procedure.

The experimental histogram of AR1 3E 30 has a substantial tail which may contribute to the large design extreme value prediction. The last four bins (out of 12) each have only one observation. A data sensitivity study was conducted in order to determine how sensitive the design values are to the tail portion of the measured data. The results of this study are presented in the next section.

#### SENSITIVITY OF RESULTS ON EXPERIMENTAL DATA

A data sensitivity study on cases AR1 3E 30, S14 3H 10, S14 3Q 10, and S82 3Q 10 was carried out in order to make an assess-

ment of how sensitive predicted results are to the experimental data. Computations were carried out for several cases for which four observations from a rather distinct tail were dropped. The results are presented in Table 1.

The results indicate that elimination of a small percentage of the experimental data can result in large changes in the design extreme value, particularly when  $T > -0.5$  and  $c < 0.0$ . Based on these results, it appears that a substantially large number of observations in the tail portion of the experimental data is desirable in order to predict results that are insensitive to minor errors in the observations.

It should be noted that the data sensitivity study as conducted above is rather severe, i.e., four points were removed in a systematic manner from the tail portion of the measured histogram. The probability that four points from the tail would be incorrectly measured is rather small. One way of determining how long an experiment should be conducted and consequently how many observations should be made is to monitor various statistics such as moments during any particular experiment. Once these statistical measures reach steady state, the sample can be judged adequate. Proceeding in this manner may have been impractical during a single prototype run. However, combining several runs under nearly identical experimental conditions may prove possible.

TABLE 1  
RESULTS OF SENSITIVITY  
STUDY ON FOUR REPRESENTATIVE CASES

CASE	$m$	$c$	$\lambda$	$X_s$	$X_D$	$X_D/X_s$	$N$	$\% \Delta X_D$	$\% \Delta X_s$
AR1 3E 30	105. 23.	-0.21 0.50	$0.2 \times 10^{-10}$ 123.3	8.19 6.96	29.16 16.80	3.56 2.41	60 56	-42	-15%
S14 3H 10	11.9 7.8	0.33 0.42	209.0 13.9	25.12 23.15	174.0 137.8	6.93 5.95	209 205	-21	-8%
S14 3Q 10	9.7 26.3	-0.37 -0.24	$0.1 \times 10^{-2}$ $0.5 \times 10^{-6}$	8.18 6.77	155.0 80.4	19.0 11.9	131 127	-93	-17%
S82 3Q 10	11.8 42.8	-0.31 -0.18	$0.3 \times 10^{-3}$ $0.6 \times 10^{-9}$	5.0 3.6	85.3 35.2	17.1 9.8	63 59	-59	-28%

## EVALUATION OF THE MAXIMUM LIKELIHOOD METHOD

The method previously employed to determine estimation parameters for the generalized gamma distribution was based on the Stacy-Mihram method (see Reference 1 and 2). One of the advantages of the Stacy-Mihram method is that negative values of the parameter  $c$  are allowed. However, based on the current analysis of over 100 cases, it appears that cases for which  $c$  is negative result in unrealistically large values of the ratio of the design extreme value to the significant value regardless of the processing technique (Stacy-Mihram method, grid search, nonlinear least squares method). Consequently, development of representations of the experimental via the generalized gamma function should be restricted to cases for which the parameter  $c > 0$ . More detailed analysis (see comments at the beginning of this section) has revealed that processing should be restricted to those cases for which the normalized logarithmic moment is less than  $-0.5^*$ . If the above restrictions are imposed, the preferred processing method is the maximum likelihood method as developed in a previous section of this report. One of the reasons the maximum likelihood method is preferred is that it provides unbiased and efficient parameter estimates in the asymptotic limit, i.e., as the number of samples becomes

\*Examination of the data shows that it cannot be adequately represented by any of the standard distributions.

large, provided specified regularity conditions hold (see References 6 and 7 for details). Also, the method is applied to the actual experimental measurements, i.e., grouping the data into histograms is not necessary.

A comparison of the maximum likelihood results with experimental data is presented in Figure 2. As demonstrated by the figure, the agreement between the generalized gamma representation and the experimental data is excellent. The ratio of design extreme value to most significant value is 3.65 which is less than the value obtained by the Stacy-Mihram method is 5.58.

#### DIGITAL PROGRAM DESCRIPTION

The current digital program is based on a previous program that implemented the Stacy and Mihram method. The following modifications have been made:

1. The maximum likelihood method has been incorporated into the previous program. This is accomplished by replacing subroutine GETCML with a new version. The maximum likelihood method should be used only when the numbers of observations is greater than 100, or the number of bins is greater than 60. The Stacy-Mihram method should be used when the number of bins is less than 20 and the number of observations is greater than 100.

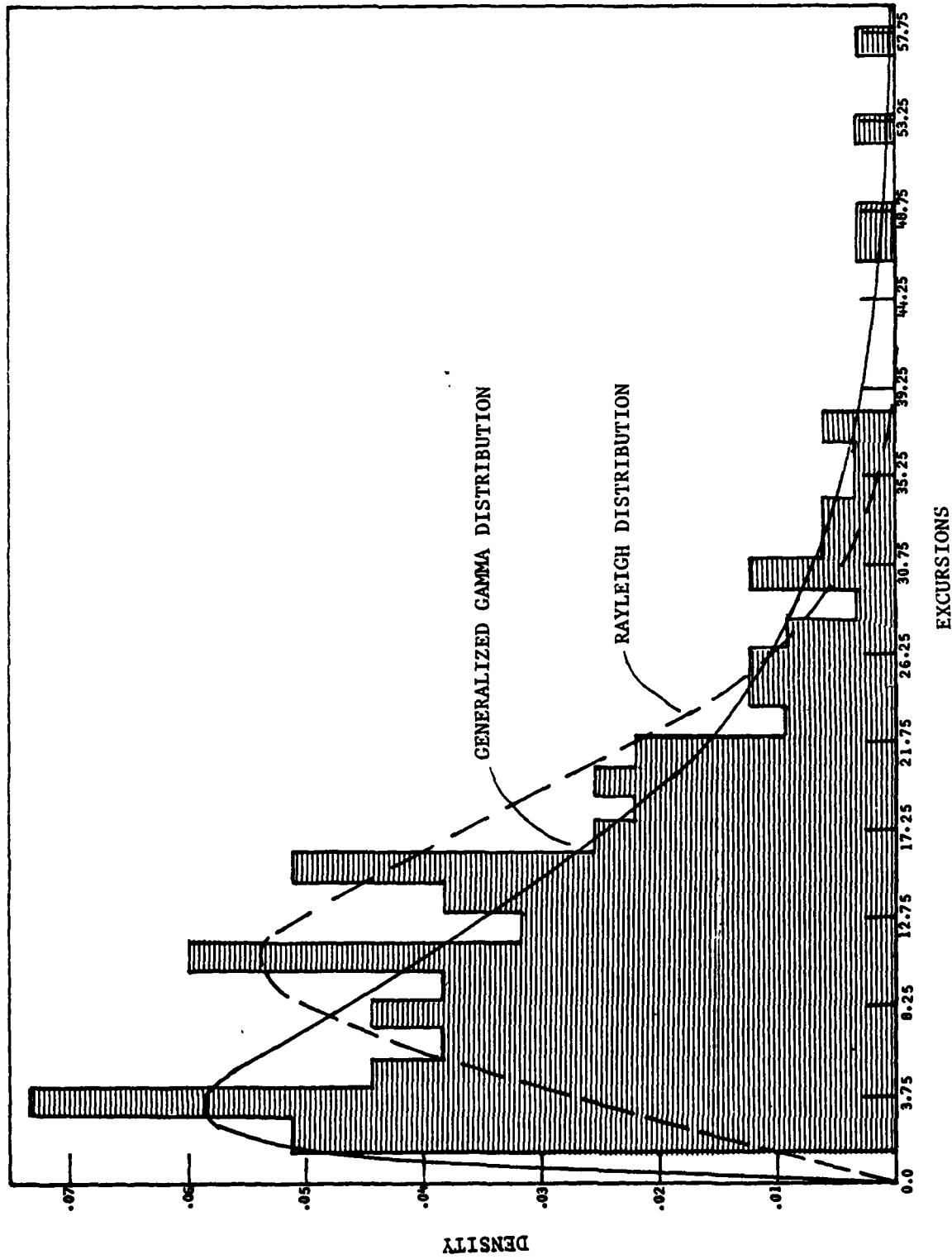


FIGURE 2. A COMPARISON OF THE GENERALIZED GAMMA DISTRIBUTION AS DETERMINED BY THE MAXIMUM LIKELIHOOD METHOD WITH EXPERIMENTAL DATA

2. As an added option, data can be read in the form of densities or counts per bin.
3. Printer plots of Rayleigh, generalized gamma function, and experimental data are generated.
4. Since the extreme values are a function of the number of observations (time), time effect plots are generated along with tabular results.

Three proprietary programs supplied by DTNSRDC are used in the present system. These are PLOTPR, MGAMMA, and MDGAM. PLOTPR is used to produce printer plots of data written on unit 3, MGAMMA computes the gamma function, and MDGAM computes the incomplete gamma function. Printer plot programs and programs for the gamma function are readily available on most computer systems. Hence, replacing these with programs available at different installations should present no major difficulties. The incomplete gamma function can be obtained from the gamma function via a simple integration procedure. However, due care must be exercised to assure sufficient numerical accuracy.

A complete description of input data, printed output, and subroutines follows. A listing of the program is presented in Appendix A. An example of print out is presented in Appendix B.

#### INPUT DATA

A description of input data follows. Card images of a typical run stream are illustrated in Figure 3.

<u>Card</u>	<u>Variable</u>	<u>Cols.</u>	<u>Format</u>	<u>Description</u>
1	IC(1)	1-2	I2	option to convert input data from frequencies to densities. $H_{NEW} = H_{OLD}/(N \times DEL)$ ON: IC(1) = 1
2	IC(2)	3-4	I2	option to convert L values from time to sample/run time $L_{NEW} = (L_{OLD}/Run time) \times N$ ON: IC(2) = 1
3	IC(3)	5-6	I2	option to change random variable values to midpoint values ON: IC(3) = 1
1	IC(4)	7-8	I2	option to convert input data by factor CAL ON: IC(4) = 1
1	IC(5)	9-10	I2	option to plot Histogram, Rayleigh and Gamma with "PLOTPR" ON: IC(5) = 1
1	IC(6)	11-12	I2	option to plot above in MKS units (change by factor CONVER) ON: IC(6) = 1
1	IC(7)	13-14	I2	option to plot L in hours vs. $X_D$ , $X_G$ (Time effect plots) ON: IC(7) = 1
1	IC(8)	15-16	I2	option to plot above in MKS units ON: IC(8) = 1
1	IC(9)	17-18	I2	option to print gamma functions for all L values OFF: IC(9) = 1 ON: otherwise

<u>Card</u>	<u>Variable</u>	<u>Cols.</u>	<u>Format</u>	<u>Description</u>
1	IC(10)	19-20	I2	option to print above in MKS units ON: IC(10) = 1
1	IC(11)	21-22	I2	scale on Time Effect Plots Each dot will be spaced by 10 ** (IC(11)). -9 < IC(11) < 99 Default: .05, 1-5 hrs.
1	IC(12)	23-24	I2	IC(12)=0 calculates design values and option for plots. IC(12)=1 no design values calculated and no plots printed IC(12)=2 no design values calculated, but option for plots.
1	IC(15)	29-30	I2	debug printout for Stacy-Mihram method
1	IC(16)	31-32	I2	number of double pages on plot Default: 1 double page minimum is 1 double page (2 pages)
1	IC(20)	39-40	I2	IC(20)=1: Use maximum likelihood method in estimating parameters Otherwise: Use Stacy-Mihram method
1	ISTOPD	41-50	I10	Number of increments in excess of 200 to be used in determining various statistical measures. Recommended value is zero or one.
1	IPLOT	51-55	I5	IF IPLOT # 0, Calcomp plots will be generated
1	IRAY	56-60	I5	IF IRAY # 0, Rayleigh distribution will be generated

<u>Card</u>	<u>Variable</u>	<u>Cols.</u>	<u>Format</u>	<u>Description</u>
1	CONVER	61-70	F10.0	Conversion factor to go from English to MKS units
1	CAL	71-80	F10.0	Conversion factor of input data to measurement desired
2	IOP(1)	1-2	I2	IOP(1)=0: $c = 2.0$ as initial guess for use in Wegstein iteration IOP(1)=1: Initial $c$ value read in.
2	IOP(2)	3-4	I2	IOP(2)=1: Write initial $c$ value
2	IOP(3)	5-6	I2	IOP(3)=1: Write $c, m, \lambda$ estimates obtained by maximum likelihood estimation
2	IOP(5)	9-10	I2	IOP(5)=1: Debug printout
3	N	1-5	I5	Total number of samples
3	K	6-10	I5	Number of divisions
3	ALP	11-20	F10.0	Value of $i - F(x)$ corresponding to the desired extreme value
3	DEL	21-30	F10.0	Division size
3	CMNT(1-5)	31-80	5A10	Description of the set of data
4	NL	1-10	I10	Number of L values to be considered. Up to five can be used (See Equation 2.15 in Reference 2.)
4	ELL(1-5)	11-60	5F10.0	up to five values of L
5	X(1), H(1)	1-80	8F10.0	K pairs of ordinate and experimental density values to be analyzed 4 pairs per card.

Multiple cases are processed by stacking the data sets with the last card as a blank to serve as a flag indicating the last case. However, the option cards (first two cards) are only read once. A run consisting of two sets of data is illustrated in Figure 3.

XXXXXX,CM77003,T200,F4.  
CHARGE,TTTT,NNNNNNNNNN,CC,K.  
ATTACH,BINAR,BBBBBBBBBB,1D=TTTT.  
ATTACH,NSRDC.  
ATTACH,IMSL.  
LOAD(LIB=IMSL/NSRDC)  
BINAR.

8 1 1  
1 1  
311 15 0.01 12.0 P22 H 30  
5 311. 833. 7492. 4998. 9996.  
6. .0075 18. .0147 30. .0212 42. .0174  
54. .0102 66. .0151 78. .0079 98. .0016  
102. .0005 114. .0005 126. .0005 138. .0  
150. .0 162. .0 174. .0007  
262 14 0.01 3.0 P22 3F 10  
5 262. 546. 1638. 3274. 5550.  
1.5 .01527 4.5 .03817 7.5 .05089 10.5 .04962  
13.5 .05598 16.5 .03435 19.5 .04071 22.5 .01908  
25.5 .01272 28.5 .00763 31.5 .00382 34.5 .00127  
37.5 .00127 40.5 .00254

BLANK CARD

8  
8

FIGURE 3. CARD IMAGES FOR A TYPICAL COMPUTER RUN

## PRINTED OUTPUT

The printed output is dependent upon the options that are chosen. A general description of the printout follows. A sample run is shown in Appendix B.

The first page consists of the user options chosen. The options are only printed once for all the data sets. The initial guess for  $c$ , the value of  $c$  at each iteration used in finding  $c$  and the estimated parameters are next, if maximum likelihood method is chosen. If Stacy-Mihram method is chosen the iterations for finding  $m$  are given.

The second page identifies the data set being processed. The number of observations, number of bins,  $\alpha$ , and bin size are next printed followed by the bin widths, the histogram data. Next the logarithmic moments of the data follows;  $\bar{y}$ ,  $s$ , and  $T$ . The last items in the header are the calculated generalized gamma parameters  $m$ ,  $c$ ,  $\lambda$  as computed by the maximum likelihood method or Stacy-Mihram method.

Eight columns follow: the values of the random variable, the value of the random variable in MKS units if CONVER is given and Option 6 of the IC options is chosen, the corresponding Rayleigh predicted frequencies, the corresponding measured densities, the corresponding Rayleigh predicted densities and the corresponding generalized gamma predicted densities.

The values of  $L$ , the density function, the cumulative distributions, and corresponding random variables are printed for each value

of L. These occupy eight columns of output arranged as follows:

Column 1 - running index

Column 2 - value of the random variable.

Column 3 -  $f(x)$ , (Equation 1 in Reference 1)

Column 4 -  $f(x)$ , (Equation 11 in Reference 1)

Column 5 -  $f(x)^L$

Column 6 -  $G(x)$ , (Equation 12 in Reference 1)

Column 7 -  $h(x)$ , (Equation 14 in Reference 1)

Column 8 -  $H(x)$ , (Equation 15 in Reference 1)

The following page gives the Rayleigh design values for all L values.

The fourth page gives the generalized gamma design values for one L value. These include  $x_M$ ,  $x_0$ ,  $x_S$ ,  $x_G$ ,  $x_{GA}$ ,  $x_D$ , and  $x_{DA}$ . To the right of these values,  $x_0$ ,  $x_G$ ,  $x_{GA}$ ,  $x_D$ , and  $x_{DA}$  as computed by a table look-up procedure are printed. This page is repeated for each L value.

After these pages there is a comparison of design values for all L values. The next pages consist of a printer plot which show the measured density, the density as predicted by the generalized gamma function, and the density predicted by the Rayleigh density function.

The following pages consist of a time effect printer plot which give L in hours vs.  $x_D$  and  $x_G$ . The CalComp plots generated replicate the printer plots.

## PROGRAM DESCRIPTIONS

### Main Program

The main program serves as a driver for the entire digital program. In addition, many of the required computations are performed in the main program. A description of program flow follows.

All data required for the program is read from unit 5. A description of this data is presented earlier in connection with the input data.

After reading the data, the moments of the logarithms of the random variables are computed along with the Rayleigh parameter and maximum histogram value. The normalized third logarithmic moment,  $T$ , is tested against the value -0.5. If  $T > -0.5$ , the case is not processed. The parameters  $m$ ,  $c$ , and  $\lambda$  are computed by a call to subprogram GETCML via the maximum likelihood or Stacy-Mihram methods.

Once values for  $m$ ,  $c$ , and  $\lambda$  are determined, the various statistical measures are computed for up to five values of  $L$ . The density functions and cumulative distributions are printed for each  $L$  value considered, followed by the extreme, most probable, and significant values. The extreme, most probable, and significant values are obtained by the Wegstein iterative method (see Subroutine WEG).

After the above computations are performed, CalComp and printer plots are generated. These consist of comparisons of the measured density function with the generalized Gamma density function and with the Rayleigh density function.

Descriptions of major FORTRAN symbols are presented in Table 2.

TABLE 2  
MAJOR FORTRAN SYMBOLS USED IN MAIN PROGRAM

<u>FORTRAN Symbol</u>	<u>Math Symbol</u>	<u>Description</u>
N	N	total number of samples
K	K	number of divisions
ALP	$\alpha$	value of $1-F(x)$ corresponding to the desired extreme value
DEL	$\Delta x$	bin size
NL	-	number of L values
ELL	L	array of L values
ISTOPD	-	number of increments in excess of 200 to be used in determining extreme values
X	x	array of ordinate values
H	$f_i$	experimental density values
S	S	value related to second logarithmic moment
YBAR	$\bar{y}$	first logarithmic moment
T	$\bar{T}$	value related to third logarithmic moment
R	R	Rayleigh parameter
C	c	parameter c
M	m	parameter m, real
AMBDAA	$\lambda$	parameter lambda
XO	x	ordinate value
FOXO	$f(x; \lambda, m, c)$	predicted density function
BFOXO	$F(x)$	cumulative distribution function
BFOXL	$F(x)^L$	cumulative distribution raised to the Lth power
GOFX	$g(x)$	density function corresponding to $F(x)^L$
HLO	$h(x)$	asymptotic density function
HHO	$H(x)$	asymptotic cumulative distribution
XSUBM*	$x_\mu$	most probable value
XSUBO*	$x_0$	value at which $F(x) = 2/3$
XSUBS*	$x_s$	significant value
XSUBG*	$x_g$	most probable extreme value

\*These symbols are used as labels in the printed output generated by the program.

TABLE 2  
MAJOR FORTRAN SYMBOLS USED IN MAIN PROGRAM (Cont.)

<u>FORTRAN Symbol</u>	<u>Math Symbol</u>	<u>Description</u>
XSUBGA*	$x_{ga}$	most probable asymptotic value
XSUBD*	$x_D$	design extreme value
XSUBDA*	$x_{da}$	extreme asymptotic value

\* These symbols are used as labels in the printed output generated by the program.

### Subroutine GETCML

Purpose: To determine the parameters of a generalized gamma distribution from statistical measures of input data by either Stacy-Mihram method or maximum likelihood method.

Usage: CALL GETCML(AT,U,E,S,YBAR,GAMMA,OOC,T)

COMMON/PARAMS/,/PRNT/,/HIST/,/INPUT/,/OPT/,/IOUNIT/

#### Subroutines and Subprograms Called:

MLHCM1	Maximum likelihood estimation
WEG	Wegstein iteration
PSI	Digamma function
PSI1	Trigamma function
DIST	Evaluates measured and observed frequencies and densities
MGAMMA	Gamma function

Description of Parameters: See Table 3.

#### Remarks:

- (1) If  $X_D$  does not converge, a check value is sent to calling routine.

TABLE 3  
MAJOR FORTRAN SYMBOLS

<u>Name</u>	<u>Location</u>	<u>Description</u>
AMBDA	/PARAMS/	$\lambda$ parameter of output distribution
C	/PARAMS/	c parameter of output distribution
M	/PARAMS/	m parameter of output distribution
X(I)	/INPUT/	random variable values
ICK1	/OPT/	Check value. If ICK1=1, c root was not found via maximum likelihood method
IOP(I)	/OPT/	options
YBAR	argument	logarithmic mean of input data
S	argument	second log moment of input data
T	argument	third log moment of data
CMNT(I)	/INPUT/	header label for GETCML output
FOBS, FTH	/PRNT/	computed values for observed and theoretical frequencies
DENO, DENT	/PRNT/	computed values for observed and theoretical densities
OOC	argument	reciprocal of c
ICK	argument	check value. If $X_D$ does not converge, ICK=1 is sent back to calling routine.
U	argument	trigamma function of final m value
E	argument	digamma function of final m value

TABLE 3  
MAJOR FORTRAN SYMBOLS (Cont.)

<u>Name</u>	<u>Location</u>	<u>Description</u>
AT	argument	absolute value of the third log moment of data
H(I)	/HIST/	densities corresponding to random variable values.

### Subroutine MLHCM1

Purpose: To estimate parameters  $\lambda$ ,  $c$ , and  $m$  by the maximum likelihood method for the generalized gamma function.

Usage: CALL MLHCM1

COMMON/PARAMS/,/HIST/,/INPUT/,/IOUNIT/,/PRNT/,/OPT/

#### Subroutines and Subprograms Called:

WEG	Wegstein iteration
PSI	Digamma function
F	Function of parameter $c$
G	Function used by WEG to determine $c$

#### Description of Parameters:

<u>Name</u>	<u>Location</u>	<u>Description</u>
AMBDA	/PARAMS/	$\lambda$ parameter of output distribution
C	/PARAMS/	$c$ parameter of output distribution
M	/PARAMS/	$m$ parameter of output distribution
X(I)	/INPUT/	random variable values
O(I)	/HIST/	densities corresponding to random variable values
IOP(I)	/OPT/	options
ICK1	/OPT/	check value. If ICK1=1, $c$ root was not found

#### Remarks:

- (1)  $c$  is restricted to values greater than zero.

- (2) If a root for  $F(c)$  is not found, a check value is sent to the calling routine.
- (3) An initial guess for  $c$  for use in WEG may be read in. Otherwise a default value of 2.0 is used.

## Subroutine DIST

Purpose: To calculate frequency and density distributions for Rayleigh and generalized gamma distributions.

Usage: CALL DIST (all parameters and arguments are passed in COMMON Blocks)

### Subroutines and Subprograms called:

MDGAM evaluates incomplete Gamma function

### Description of Variables:

<u>Variable</u>	<u>Location</u>	<u>Description</u>
X(I)	/INPUT/	random variables values
H(I)	/INPUT/	array of histogram ordinates for input data
K	/HIST/	number of bins in input histogram
N	/INPUT/	number of observations in input histogram
FTH(I), FOBS(I)	/PRNT/	computed values for the observed frequency distribution and theoretical frequency distributions
DENO(I), DENTR(I)	/PRNT/	computed values for the observed density distribution and theoretical frequency distributions

Remarks: The input histogram need not be evenly spaced in the abscissa x.

Method: Generalized gamma cumulative distribution is evaluated by MDGAM. The density function is obtained by numerically differentiating the obtained cumulative. The Rayleigh density distribution is obtained in a like manner.

### Subroutine WEG

Purpose: To determine root of  $x = f(x)$  by Wegstein iteration.

Usage: CALL WEG (I, X, J, N)

Description of parameters:

I - input, iteration count. Initialization action is taken when I = 1.

X - input, estimate of root of  $x = f(x)$ . On output, refined estimate of root.

J - number of significant digits of accuracy desired in x for solution -- used only in initialization.

N - output - completion flag

0 convergence not obtained

1 convergence is within given accuracy

Remarks: The function whose root is to be found is computed externally. The calling sequence is as follows.

(1) Initialize by issuing CALL WEG (1, XG, J, N) where XG is the initial guess for the root and J is the tolerance described above. N is irrelevant.

(2) Set up a loop to calculate  $XN=F(X)$  using output X from WEG

(3) The simplest calling sequence would then be

```
X=XG           Initialize X to guess value
DO 1  I=1,ITERCT
      CALL WEG(I,X,ITOL,N)  First pass will initialize WEG
      IF (N.EQ.1) GO TO 2
      X = F(X)
1  CONTINUE
C  MAXIMUM NUMBER OF ITERATIONS REACHED WITHOUT CONVERGENCE
STOP
2  CONTINUE
C  ROOT FOUND WITHIN TOLERANCE
```

Method: The Wegstein method refines the root of the equation  $x = f(x)$  by calculating an improved estimate  $x_{I+1}$  which is the intersection point of the line  $y = x$  and the secant line of  $f(x)$  based on the previous two evaluations of  $f(x)$ . This method requires only one function evaluation per iteration.

The equation for the intersection point is given by

$$x_{I+1} = \frac{F(x_I) * x_{I-1} - F(x_{I-1}) * x_I}{F(x_I) - F(x_{I-1}) - x_I + x_{I-1}}$$

Completion code N is set 1 when

$$|x_{I+1} - x_I| < |10^J * x_{I+1}|,$$

i.e., when the change in  $x$  between iterations is less than one part in  $10^J$ .

#### Function PSI1

Purpose: To evaluate the trigamma function,  $\psi'(x)$

Usage: PSI1(x)

#### Description of Parameters:

PSI1(x) - value of the trigamma function at argument

x (Output)

x - argument of the function (Input)

#### Remarks:

(1) Argument x must be greater than zero.

(2) The trigamma function is the second derivative of the natural logarithm of the gamma function.

Method: For arguments greater than 13 an asymptotic expansion,

$$\psi'(x) \simeq \frac{1}{x} + \frac{1}{2x^2} + \frac{1}{6x^3} - \frac{1}{30x^5} + \frac{1}{42x^7} - \frac{1}{30x^9}$$

is used. The truncation error associated with the above expansion is less than  $1 \times 10^{-10}$  for  $x > 13$ . For arguments between 0 and 13, the recursion relation

$$\psi'(x) = \psi'(m+x) + \sum_{i=1}^x \frac{1}{(x+i-1)^2}$$

is used, where  $m$  is chosen as the smallest integer for which  $x + m \geq 13$ .  $\psi'(m+x)$  is evaluated by the previous formula. The formulas are obtained from Reference 8.

#### Function PSI2

Purpose: To evaluate the tetragamma function,  $\psi''(x)$

Usage: PSI2(x)

Description of Parameters:

PSI2(x) - value of tetragamma function at argument X  
(Output).

X - argument of the function (Input).

Remarks:

- (1) argument must be greater than zero.
- (2) the tetragamma function is the third derivative of the natural logarithm of the gamma function.

Method: For arguments greater than 13, the asymptotic expansion,

$$\Psi''(x) \approx \frac{1}{x^2} - \frac{1}{x^3} - \frac{1}{2x^4} - \frac{1}{6x^5} + \frac{1}{6x^6} - \frac{1}{10x^7} - \frac{1}{6x^8}$$

is used. The truncation error associated with the above expansion is less than  $1 \times 10^{-10}$  for  $x > 13$ .

For arguments between zero and 13, the recursion relation

$$\Psi''(x) = \Psi''(x+m) - \sum_{i=1}^k \frac{1}{(x+i-1)^3}$$

is used, where  $m$  is the smallest integer which satisfies  $x + m \geq 13$ , and  $\Psi''(x+m)$  is evaluated by the previous formula.

The above formulas are obtained from Reference 8.

#### Function PSI

Purpose: To evaluate the digamma function,  $\Psi(x)$ .

Usage: PSI(x)

#### Description of Parameters:

PSI(x) - value of digamma function at argument x (Output)

x - argument of function (Input)

#### Remarks:

(1) Argument X must be greater than zero.

(2) The digamma function is the derivative of the natural logarithm of the gamma function.

Method: For arguments greater than 13 an asymptotic expansion is used:

$$\Psi(x) \approx \ln x - \frac{1}{2x} + \frac{1}{12x^2} + \frac{1}{120x^4} - \frac{1}{252x^6}$$

The truncation error associated with the above expansion is less than  $1 \times 10^{-10}$  for  $x > 13$ . For arguments less than 13, the recursion relation

$$\Psi(x) = \Psi(x+m) - \sum_{i=1}^k \frac{1}{x+i-1}$$

is used, where  $m$  is the smallest integer which satisfies  $x + m \geq 13$ , and  $\Psi(x + m)$  is evaluated by the previous formula. The above formulas are obtained from Reference 8.

#### Subroutine MGAMMA

Purpose: To evaluate Gamma function

Usage: CALL MGAMMA (X, GAMMA, IER)

#### Description of Parameters:

X - input, argument of Gamma function

GAMMA - output value of Gamma function at X

IER - error code

Remarks: This is a library routine supplied by DTNSRDC.

User must attach library (IMSL).

Error code meaning is unknown.

### Subroutine MDGAM

Purpose: To evaluate incomplete gamma function

Usage: Call MDGAM(T, X, PR, IE)

Description of Parameters:

T input, limit of integral

X input, exponent in integral

PR output, value of incomplete gamma function

IE output, error code

Remarks:

- (1) This is a library provided by DTNSRDC
- (2) User must attach IMSL library
- (3) Meaning of error code unknown
- (4) Function evaluated is given by,

$$P(x, T) = \frac{1}{\Gamma(x)} \int_0^T u^{x-1} e^{-u} du$$

### Subroutine HISTO

Purpose: To set up labels and plot the experimentally measured density function.

Usage: Call HISTO

Subprograms and function called: SCALE, AXIS, SYMBOL, LINE

Remarks: This program takes the measured density, which is passed in COMMON and uses these values to generate data for a CalComp plot.

## Function G

Purpose: To be used by WEG in evaluating the c parameter by the maximum likelihood method.

Usage: G(c)

Subroutines and Subprograms Called: PSI

Description of Parameters:

G(c) - value of function (Output)

c - argument of the function (Input)

Remarks: c must be greater than zero.

Method: To implement the Wegstein technique the equation  $f(x) = 0$  must be put into the form  $x = g(x)$ . The equation used is given by

$$\begin{aligned} c = & \frac{\sum_{i=1}^k o_i}{\sum_{i=1}^k o_i \ln x_i} \left[ \ln c + \ln \left( \sum_{i=1}^k x_i^c o_i \right) \right. \\ & \left. + \ln \left[ \frac{\sum_{i=1}^k \ln x_i (x_i^c o_i)}{\sum_{i=1}^k x_i^c o_i} - \frac{\sum_{i=1}^k o_i \ln x_i}{\sum_{i=1}^k o_i} \right] - \ln \left( \sum_{i=1}^k o_i \right) \right] \\ & + \psi \left[ \frac{1}{c \left( \frac{\sum_{i=1}^k \ln x_i (x_i^c o_i)}{\sum_{i=1}^k x_i^c o_i} - \frac{\sum_{i=1}^k o_i \ln x_i}{\sum_{i=1}^k o_i} \right)} \right] \end{aligned}$$

$o_i$  is the density

$x_i$  is the value of random variable

## Function F

Purpose: To evaluate the function used to find the c parameter by the maximum likelihood method.

Usage: F(c)

Subroutines and Subprograms Called: PSI

Description of Parameters:

F(c) - value of the function used to find the parameter c (Output)

c - argument of the function (Input)

Remarks: c must be greater than zero

Method: The equation used is given by

$$\begin{aligned} F(c) &= \ln c + \ln \left( \sum_{i=1}^k x_i^c o_i \right) \\ &+ \ln \left[ \frac{\sum_{i=1}^k \ln x_i (x_i^c o_i)}{\sum_{i=1}^k x_i^c o_i} - \frac{\sum_{i=1}^k o_i \ln x_i}{\sum_{i=1}^k o_i} \right] - \ln \sum_{i=1}^k o_i \\ &+ \psi \left[ \frac{1}{c \left( \frac{\sum_{i=1}^k \ln x_i (x_i^c o_i)}{\sum_{i=1}^k x_i^c o_i} - \frac{\sum_{i=1}^k o_i \ln x_i}{\sum_{i=1}^k o_i} \right)} \right] \\ &- \frac{c \sum_{i=1}^k o_i \ln x_i}{\sum_{i=1}^k o_i} = 0 \end{aligned}$$

$o_i$  is the density

$x_i$  is the value of the random variable

## CONCLUSIONS

An extensive study of an existing digital computer program for determining parameters required to represent SES 100B trials data by a generalized gamma function was carried out. Difficulties encountered in earlier studies in which approximately 5 percent of the cases analyzed produced unrealistic values for the distribution parameters have been resolved. In order to resolve this problem, data sensitivity studies were carried out and three other techniques besides the Stacy-Mihram method were investigated. Based on results of the study, criteria are developed for determining cases that should not be processed. A summary of major conclusions follows.

- a. The Stacy-Mihram, maximum likelihood, grid search, and nonlinear least squares methods all yield estimated probability distribution functions which represent very well the experimentally measured histograms.
- b. Design extreme values are more sensitive to the experimental measurements than to the numerical procedure used in determining the estimation parameters.
- c. In some cases the predicted extreme design values are unrealistically large whereas the significant values are realistic. Based on an analysis of over 100 cases, it is concluded that cases having a nondimensional third logarithmic moment greater than -0.5 produce unrealistic extreme design values. Results of a data sensitivity

study have revealed that in all these cases the data may not be considered as a sample taken from a steady-state random phenomenon, and hence it is recommended that statistical analyses for these cases not be carried out.

- d. The Stacy-Mihram method should be used when the number of bins (number of divisions used in the analyses) is small (< 20) and when the number of samples is large (> 100).
- e. The maximum likelihood method should be used when the number of bins is large (> 60) or when the sample size is large (> 100).

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APPENDIX A  
COMPUTER PROGRAM LISTING

```

PROGRAM PR1K(INPUT,OUTPUT,TAPES=INPUT,TAPPE=OUTPUT,TAPE7=TAPE3)
REAL MAXY
REAL V
DIMENSION XM(500),Z(500),FM(500),PM(500),IVAR(4)
DIMENSION XMS(500),PMMS(500),FMMS(500),ZM(500)
DIMENSION XOMS(200),PMOMS(200),ZOMS(200)
*30FXM(200),ZL0MS(200),HHSMS(200)
DIMENSION PDX(7),ELT(7),FSC(7),PDXM(7),FGX(7),RGXM(7)
COMMON/I*PUT/X(100),ELL(7),ALP,0(100),ISTOPP,IPLOT,IPAY,IC(20)
COMMON/OPR/ICP(20),ICK1
COMMON/PLO/NPUN,NPLCT,ITP(4),ITY(5),ITX(5),
SHMAG,MAXSCA,SCA(10),FPM(10)
COMMON/PNT/EPBS(100),FT4(100),DEN(100),DELT(100)
S,FTHP(100),UFNT(100),CONVER,CAL
COMMON/HIST2/FIRSTX,DELY,FIRSTY,DELY
COMMON/HIST/PINLO,K,PIV(100),H(100),YMAX,XLEN,YLEN,CNNT(5)
COMMON/PARAMS/V,AMBDA,C,P
COMMON/TCUNIT/ICP,IRD,IPRT,IPCH
DATA ICB,IPD,IPRT,IPCH/6.5.6.6/
DATA XLEN,YLEN/10.,0./
TTH=2./3.
CALL PLTS(0,3,7)
READ OPTIONS
READ(5,1F) (TC(I),I=1,20),ISTOPP,IPLOT,IPAY,CONVER,CAL
1413 FORMAT(14,3H IC OPTIONS USED IN THIS RUN SET IPLOT=,I*,*,Y,
*EIPAY=,T*)
DO 14 I=1,20
14 IF(IC(I).GT.0) WRITE(6,1414) I,IC(I)
1414 FORMAT(14,2H OPTION ,I2.14=,I2)
WRITE(6,2013)
2013 FORMAT(1H0,3HICP OPTIONS USED IN THIS RUN SET)
READ(5,115) (ICP(I),I=1,20)
115 FORMAT(20I2)
DO 979 I=1,20
IF(ICP(I).GT.0) WRITE(6,1414) (I,ICP(I))
979 CONTINUE
IF(CONVER.EQ.0.) CONVER=1.0
15 FORMAT(20I2,7I10+2I5,2F10.0)
C
99 CONTINUE
VV(1)=0.
FM(1)=0.
TP=1
PM(1)=0.
KLL=XLEN+2
C
***** READ INPUT AND PRINT *****
C
READ (5,100) "K,ALP,DEL,CNNT"
100 IF(K.EQ.0) GO TO 10003
FORMAT(2I5,2F10.0,5A10)
C
C READ L VALUES
READ(5,102) AL,(ELL(I),I=1,UL)
102 FORMAT(I10,7F10.0)
READ(5,101) (Y(I),H(I),I=1,K)

```

```

101 FORMAT( PF10.0 )
200 FORMAT( 1X, 2F10.4 )
  OMA = 1.-ALP

C
C   IF UNEQUAL INTERVALS DEL=?
C   DEL READ IN FROM X ARRAY
C   IF STARTING POINT N.E. 0,4 - SIGN WILL PRECEDE
C   THE STARTING VALUE GIVEN AS DEL.
  DXY=DEL
  IF(DEL.GT.0.0) GO TO 103
  ST=-DEL
  D(1)=X(1)
  DSUM=D(1)
  X(1)=ST
  DMX=D(1)
  DO 196 I=2,M
    D(I)=X(I)
    IF(C(I).GE.D(I-1)) D"X=D(I)
    X(I)=X(I-1)+C(I-1)
    DSUM=D(I)
    DSUM=DSUM+C(I)
196  CONTINUE
  IF(IC(3).EQ.0) TFLG=1
  IF(TFLG.EQ.1) GO TO 1000
  GO TO 107
109  DO 109 I=1,K
109  D(I)=DEL
107  CONTINUE
C   OPTION 1: CONVERT INPUT TO DENSITY IF NECESSARY
  IF(IC(1).NE.1) GO TO 2
  DO 11  I=1,K
11   H(I) = H(I) / (FLOAT(N)+C(I))
C
C   OPTION 2: MODIFY L VALUES IF NECESSARY
2   IF(IC(2).NE.1) GO TO 3
  RTIME = ELL(1)
  DO 12 I=1,NL
    ELT(I) = ELL(I)
12   ELL(I) = (ELL(I)/RTIME)*"
C
C   OPTION 3: ADJUST INPUT X(I) TO MIDPOINT OF INTERVALS IF NECESSARY
3   IF(IC(3).NE.1) GO TO 4
  DO 13 I= 1, K
13   X(I) = X(I) + D(I)/2.0
C   OPTION 4: CONVERT INPUT TO "MEASUREMENT" DESIRED
4   IF(IC(4).NE.1) GO TO 5
  DO 14 I= 1, K
    H(I) = H(I) / CAL
    D(I)=D(I)*CAL
14   X(I) = X(I) * CAL
5   CONTINUE
C   IF L VALUES ARE READ IN AS SAMPLES PER TIME INTERVAL
C   COMPUTE TIME VALUES FOR PLOT
    IF(IC(2).EQ.1) GO TO 6
    RTIME=ELL(1)/"
    DO 157 I= 1, NL
157   ELT(I) = ELL(I) * RTIME/"

```

```

6      CONTINUE
R=3.
YMAX=0.
CH=0.
S1 = 0.
S2 = 0.
S3 = 0.
C
C      ***** COMPUTE R, YI *****
C
DO 1 I=1,K
T1 =4*(I)+C(I)
IF(T1.GT.YMAX)YMAX=T1
P=R+T1*X(I)**2
YI=ALOG(Y(I))
SH=SH+T1
YI2=YI+YI
YI3=YI+YI2
S1=S1+T1*YI
S2=S2+T1*YI2
S3=S3+T1*YI3
1 CONTINUE
C
C      ***** COMPUTE YBAR, S, T *****
C
EN=1
YBAR=S1
T1=EN/(EN-1.)
T2=T1*EN/(EN-2.)
T3=S2
T4=YBAR+YBAR
T5=T4*CH
S=T1*(T3-2.*T4+T5)
T= T2*(S3-3.*YEAR*T3+YEAR*(T3+T4-T5))/T**1.5
IF(T.GT.(-.5)) WRITE(6,1237)
1237 FORMAT(1H0,45HT IS GREATER THAN -.5. GO ON TO NEXT DATA SET)
IF (T.GT.(-.5)) GO TO 99
AT=ABS(T)
C
C      ***** COMPUTE V,C,LAMRDA, AND GAMMA *****
C
ICK=1
CALL GETCM(LAT,E,S,YBAR,GAMMA,DOC,T,ICK)
IF(ICK.EQ.1) GO TO 99
IF(ICK1.EQ.1) GO TO 99
WRITE(6,P014)
9014 FORMAT(1H1,22HRAYLEIGH DESIGN VALUES)
DO P012 J=1,NL
XP0=SOR(T+ALOG(FLL(J)/SLP))
XRDMS=XRD*CONVER
WRITE(6,P013) FLL(J),XRD,XRDMS
9012 CONTINUE
9013 FORMAT(1H ,3FL= .F10.4,=X,8HXRSUB= .F10.4,5X,13HXRSUP(14X)= ,
F10.4)
IF(ICK121.EQ.1) GO TO 10004
C
C      ***** COMPUTE C,V,LAMRDA PRODUCTS FOR LATER USE *****

```

```

C
      T1 = C*Y
      T1W = T1-1.
      T2 = (C*AMRDA**T1)/GAMMA
      T3 = ABS(T2)
      CM400= T1W*000

C      ***** 10000 LOOP---OVER THE L VALUES *****
C      00 10000  IL=1,NL
C      ** INITIAL GUESSES FOR X AND X   **
C
      TX0 = 1000.
      EX0 = 0.
      EXG = 0.
      TXG = 0.
      TX0=1000.
      EX0=0.
      TXG=0.
      EXG=0.
      TXDA=1000.
      EXDA=0.
      EL = ELL(IL)
      4010 FORMAT (6,10001) EL
10001 FORMAT (1H1,3H=,F7.1)
      IF (EL .GT. 300) GO TO 0001
      EML = EXP(-EL)
      GO TO 8002
8001 EML = 0.
8002 CONTINUE
      IF (IC(9).EQ.0) WRITE(6,304)
304 FORMAT (3Y,*1*,3Y,*12Y,*LF(Y)*,12Y,*F(Y)*,10Y,*F(X)L*+10Y*
1      *G(X)*,10X,*LH(X)*,12X,*H(X)*)
700 FORMAT (///* EXTRAPOLATED VALUES */)
      ISTOP = 200+ISTOP
C      ** SET UP HISTOGRAM PARAMETERS **
      I = 2
      J=0
      B1NL0=Y(1)-D(1)/2.
      MAXX=2*D(X+XLEN)
      Y0= X(1)-D(1)
      X0D=X0+D(1)/2.
      IF (X0D .LE. .01) X0=0.
701 IRES=I/2
      RES=(FL0AT(I)/2.0)-IRES
      IF (PES.EQ.0),400,(J,LT,Y)) J=J+1
      DELL=D(J)/2.0
      DFINC=D(J)/10.0
      X0=X0+DFINC
      I = I + 1
      IF (I .GT. ISTOP) GO TO 11002
      YTO = Y0
C      ** 11020 LOOP **
C      00 11000  IF=1,E
C      ***** COMPUTE SMALL F(X), PROBABILITY DENSITY FN ****
C
      Y0 = XTO+DFINC*(IF-1)

```

```

T4=X0**T1**  

T45=(4**BDA+YC)**C  

YC = EXP(-T45)  

BFOY0 = T2*T4*YC  

C  

IF(IL.GT.1) GO TO 10002  

IF(YC.GT.YMAX) GO TO 10002  

IF(IP.EQ.500) GO TO 10002  

C  

IP=IP+1  

BFOYD=BFOY0  

YC(IP)=YC  

YT1=Y0/R  

YT2=Y0*YT1  

RNX=2.*YT1*EXP(-YT2)  

BFOYD=RNX  

IF(RNX.GT.YMAX)YMAX=RNX  

IF(BFOYD.GT.YMAX)YMAX=BFOYD  

10002 CONTINUE  

C  

***** COMPUTE CAPITAL F(X), PROBABILITY DISTRIBUTION FN **  

C  

CALL MCGAM(YT45,M,PR,TF)  

IF (IE.EQ.128.CE.1E-130) WRITE (6,501) IE  

501 FORMAT(59H ERROR IN GAMMA DISTRIBUTION FUNCTION COMPUTATION.  ERROR  

.PE.1E8)  

IF (C .LT. 0.) PR = 1.-PR  

BFOYD = PR  

C  

TTT = EL*ALOG(BFOYD)  

IF (TTT.LT.-300.) GO TO 8003  

BFOYL = BFOYD**EL  

GO TO 8004  

8003 BFOYL = 0.  

8004 CONTINUE  

C  

***** COMPUTE SMALL G(X), SMALL H(X), CAPITAL H(X) *****  

C  

IF(IC(12).EQ.2) GO TO 11000  

GOFY = EL*(BFOYL/BFOYD)*BFOYD  

FX = EL*(PR-1.)  

IF (FX.LT.-300.) GO TO 8010  

FY = EXP(FX)  

HLD = EL*FY*CFY  

HMD = EX-FML  

GO TO 8011  

8010 HLD = 0.  

HMD=-EML  

8011 CONTINUE  

IF (TF .GT. 1) GO TO 11001  

XOMS(I) = X0 + CONVER  

FOXOMS(I) = FOXD / CONVER  

BFO(I) = BFOYD  

BFOL(I) = BFOYL  

GOFXMS(I) = GOFY / CONVER  

HLOMS(I) = HLD/CONVER  

HMOYS(I) = HMD

```

```

ICON = 1
IF(IIC(9).EQ.1) GO TO 11001
WRITE(6,710) I,X0,FOX0,EOF0,PE0XL,COFX,HLD,HH0
710 FORMAT(1X,I5,7G15.6)
11001 CONTINUE
C
C      ***** INITIALIZATION *****
C
      TX=ABS(PFOXL-0MA)
      IF (TX-TXD) 9000,9001,9001
9000 TXD = TX
      END = XD
9001 CONTINUE
      IF (TXG.GE.COFX) GO TO 9002
      TXG=COFX
      END=XD
9002 CONTINUE
      TX=ABS(PFOXA-TTH)
      IF (TX-TXD) 9003,9004,9004
9003 TXD = TX
      END = XD
9004 CONTINUE
      IF (TXGA .GE. HLD) GO TO 9005
      TXGA = HLD
      FYGA = X1
9005 CONTINUE
      TX = ABS(HLD-0MA)
      IF (TX-TXA) 9006,9007,9007
9006 TXDA = TX
      EXDA = XD
9007 CONTINUE
11000 CONTINUE
      XC=XTD+DELL
      IF (PFOX0.LE..999999) GO TO 701
      GO TO 11003
11002 WRITE(6+11004) 1STOP
11004 FORMAT(1X,*,N0. 0F RECORDS ST+1X,T4)
11003 CONTINUE
      IF(IIC(12).NE.1) GO TO 17
      WRITE(6,711)
      FORMAT(//1Y,T40+13HM,*,S. SYSTEM//)
      WRITE(6,304)
      DO 16 I= 1,ICON
      WRITE(6,710) J,X0MS(J),FOX0MS(J),EOF0(J),PE0L(J),COFXM(J),
      * HLDMS(J),HFMMS(J)
16   CONTINUE
17   CONTINUE
C
      IF(IIC(12).EQ.2) GO TO 970
C
C      ***** SOLVE FOR X *****
C
      TE=(M-COC)*+COC
      XSUBY = TE/AVEDA
      XSUBMS = XSUBY + CONVER

```

```

      WRITE(6,207) XSUM,XSUMMS
207  FORMAT(2X,7HXSUM= ,F12.4,3X,12HXSUM(MS)= ,F12.4)
C
C      ***** SOLVE FOR X_ USING WEGSTEIN ITERATIVE METHOD *****
C
C
      T3=AMDA**1./EL
      TT=ALP/EL
      U = EXC
      NE=0
      CALL WEG(1,U,4,0)
C      ** START ITERATION LOOP **
33  T45 = (AMDA+U)**C
      CALL WEGAM(T45,M,PR,TE)
      IF (T .LT. 0.) PR = 1.-PR
      U = U+PR*TT
C      U=U+TT*ALOG(PR)
      CALL WEG(2,U,0,NC)
      WRITE(6,205) U,PR,TT,TT
205  FORMAT(1X,4L= ,*G20.8,EX,*PR*,G20.8,EX,*T3*,G20.8,EX,*TT*,G20.8)
      IF(U .LT. 0.) U=.0001
      NE = NE + 1
      IF (NE .LT. 50) GO TO 3000
      WRITE(6,206)
206  FORMAT(34H NO CONVERGENCE OBTAINED FOR XSUM)
      I = 0
      GO TO 32
3000  CONTINUE
      IF(NE.E2.01) GO TO 33
C      ** END ITERATION FOR X_ **
C
      32 XSUMD = U
C
C      ***** EVALUATE GAMMA DISTRIBUTION FUNCTION *****
C
C      TT=(AMDA+U)**C
C      CALL WEGAM(TT,M,PR,TE)
C      WRITE(6,702) PR
702  FORMAT(1* PR *G20.8)
C
C      ***** SOLVE FOR X_ USING WEGSTEIN ITERATIVE METHOD *****
C
C
      CA = C*A*BD
      DMC=M-1.
      CM3 = C**M-2.
      EL4 = EL-1.
      U0 = EXG
      U = U0
      NE = 0
435  CALL WEG(1,U,4,0)
C      ** START ITERATION LOOP **
      J=0
43  AU = AMDA+U
      AUC = AU**C
      IF(AUC .LT. 20.) GO TO 6001
      U=1.E-10

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```

      GO TO 6002
6001 CONTINUE
      EXPT = EXPT-AUC
      F=EL** (TR+EXPT+U**T1**)
      CALL MDSAM(AUC,M,PR,IE)
      IF (C .LT. 0.) PR = 1.-PR
      U=F/((PR+U**CM3+EXPT*(CM4+C*AUC))+T2)
5002 CONTINUE
      CALL WEG(2,U,0,NC)
C      WRITE (6,703) U,PR,F,CM3,EXPT,CM4,CM,AUC,AU,T2
      703 FORMAT (* DEBUG *5G20.2)
      IF (U .GE. 0.) GO TO 6003
      U0=U0+D(1)*2
      U = U0
      GO TO 435
5003 CONTINUE
      NE = NE + 1
      IF (NE .LT. 50) GO TO 4000
      WRITE (6,200)
200 FORMAT (74H NO CONVERGENCE OBTAINED FOR XSUB0)
      U = 0
      GO TO 437
4000 CONTINUE
      IF ("C.EQ.0") GO TO 43
      C      ** END ITERATION FOR X   **
C
C      42 XSUB0 = U
437 CONTINUE
C      ***** SOLVE FOR X USING WEGSTEIN ITERATIVE METHOD *****
C
C      U0 = X(K)
      K=K+3
      U0= X(KK)
      U = U0
      NE = 0
      CALL WEG (1,U,4,0)
C      ** START ITERATION LOOP  **
5007 AU = AMRA*U
      AUC = AU**C
      CALL MDSAM(AUC,M,PR,IE)
      IF (C .LT. 0.) PR = 1.-PR
      U = U+TTH-PR
      CALL WEG(2,U,0,NC)
C      WRITE(6,703) U,PR
      IF (U.GE.0) GO TO 5004
      U=U0+D(1)
C      U = U0
      U=1.E-10
5004 CONTINUE
      NE = NE + 1
      IF (NE.LT.50) GO TO 5005
      WRITE (6,200)
200 FORMAT (74H NO CONVERGENCE OBTAINED FOR XSUB0)
      U = 0.
      GO TO 5006

```

```

5005 CONTINUE
  IF(NGC.EQ.0) GO TO 6007
C           ** END ITERATION FOR Y **

C
6006 CONTINUE
  UMS = U * CONVER
  EXOMKS = EX0 * CONVER
  WRITE(6,310) U,UMS,EX0,EXOMKS
310  FORMAT(2X,*XSUB0= *,F12.4,3X,*XSUP0(*KS)= *,F12.4,3X*
      *TYSUB0= *,F12.4,3X*TYSUP0(*KS)= *,F12.4)

C           ***** COMPUTE SIGNIFICANT VALUE Y *****

C
C           IF (U .EQ. 0.) U=EX0
C           EMT=EM+C0C
C           CALL MEGAM(EMT,GAM,IER)
C           IF(IER.EQ.129.OR.IER.EQ.170) WRITE(6,500) IER
500  FORMAT(4EH ERROR IN GAMMA FUNCTION COMPUTATION.  ERROR= ,I8)
  AU = AMRCA*U
  AUC = AU**C
  CALL MEGAM(AUC,EMT,PP,IER)
  IF (C .LT. 0.) PR = 1.-PR
  J=3.*((1.-PR)*GAM/(GAMMA*AMRDA))
  UMS = U * CONVER
  WRITE(6,312) U,UMS
312  FORMAT(2X,*XSUB0= *,F12.4,3X,*XSUP0(*KS)= *,F12.4/)

5050 CONTINUE
  YSUBGM = XSUBG * CONVER
  RGX(1L) = XSUBG
  RGXM(1L) = YSUBGM
  EXGM = EXG * CONVER
  WRITE(6,309) YSUBGM,EXGM,EXG,EXGM
309  FORMAT(2X,7HYSUBG= ,F12.4,3X,12HYSUP0(*KS)= ,F12.4,3X*
      *8HTXSLBG= ,F12.4,3X,13HTXSUP0(*KS)= ,F12.4)

C           ***** SOLVE FOR ASYMPTOTIC VALUE OF Y USING WEGSTEIN
C           METHOD
C
C           ELOG = EL/GAMMA
C           IF (C .LT. 0.) ELOG=-EL**C
  UC = EXGA
  U = (UC*AMRDA)**C
  NC = 0
  CALL WEG(1,U,4,0)
C           ** START ITERATION LOOP **
6010 IF (U .LT. 300) GO TO 8501
  UT = 0.
  GO TO 8502
8501 UT = U**"EXP(-U)
8502 U=C**"0C+FLC3*UT
  CALL WEG(2,U,0,NC)
  IF (U.EE.0.) GOTO 6011
  U0=U0+C(1)
  U = U0
  U=1.E-10
6011 CONTINUE

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```

    NF = NE + 1
    IF (NE.LT.50) GO TO 6012
    WRITE (6,2085)
2085 FORMAT(3SH NO CONVERGENCE OBTAINED FOR XSUBGA)
    U = 0.
    WRITE(6,703) U,0.0,AMRDA,C
    GO TO 6013
6012 CONTINUE
    IF(UC.EQ.0.) GO TO 6010
C           ** END ITERATION FOR X      **
C           GA
6013 U=U+0.000/AMRDA
    UMS = U + CONVER
    EXGAM = EXGA + CONVER
    WRITE(6,6014) U,UMS,FXGA,FGA
6014 FORMAT(2X,*XSUBGA= *,F12.4,2X,*XSUBGA(MKS)= *,F12.4,2X,
    E *TXSUBGA= *,F12.4,2X,*TXSUBGA(MKS)= *,F12.4)
6015 CONTINUE
    VSUBDM = XSUBD + CONVER
    PDX(IL) = XSUBC
    PDXM(IL) = XSUBCV
    EXD = EXD + CONVER
    WRITE(6,6015) XSUBD,VSUBDM,EXD,EXDM
6015 FORMAT(2X,*XSUBD= *,F12.4,2X,*XSUBD(MKS)= *,F12.4,
    $ 3X,*TXSUBD= *,F12.4,3X,*TXSUBD(MKS)= *,F12.4)
C
C           **** SOLVE FOR ASYMPTOTIC VALUE OF X USING WGSSTET
C           METHOD
C           D      ****
C
    UU = EL+ALOG(CMA)
    UC = EXDA
    U = UU
    'E = 0
    CALL WEG(1,U,4,0)
C           ** START ITERATION LOOP  **
5040 AU = AMRDA*U
    AUC = AU+C
    CALL WOGAM(AUC,0,PR,TE)
    IF (C .LT. 0.) PR = 1.-PR
    U = U+LU
    ELPR = EL+PR
    UUU = ELPR
    IF(ELPR.LE.300.) UUU=ALOG(EXP(ELPR)-1.)
    8020 U = U-LUU
    CALL WEG(2,U,0,NC)
    IF (U.GE.0.) GO TO 6043
    UU=UU+D(1)
    U = UU
6043 CONTINUE
    NE = NE+1
    IF(NE.LT.50) GO TO 6044
    WRITE (6,2089)
2089 FORMAT(3SH NO CONVERGENCE OBTAINED FOR XSUBDA)
    U = 0.
    GO TO 6045
6044 IF(UC.EQ.0.) GO TO 6040
C           ** END ITERATION FOR X      **

```

```

C                               DA
C     UM = U + CONVER
C     EXDAM = EXDA + CONVER
6145  WRITE(6,6045) U,UM,EXDA,EXDAM
6046  FORMAT(2X, *XSUCA= *,F12.4,2X,*XSUPDA(MKS)= *,F12.4,2X,
$ *TXSUPDA= *,F12.4,2X,*TXSUPDA(MKS)= *,F12.4//)
10000 CONTINUE
        WRITE(6,6047)
6047  FORMAT(1H1,T3,9HTIME(MIN),T20,9HTIME(HRS),T75,7H VALUE,TF2,
$5HXSUBC,T67,5HXSUBG,T77,10HXSUPD(MKS),T92,10HXSUPG(MKS))
        DO 99 I=1,7L
        ESC(I) = ELT(I) / 60.0
        WRITE(6,6048) ELT(I),ESC(I),ELL(I),RDX(I),PGX(I),RDXM(I),PGXM(I)
98  CONTINUE
6048  FORMAT(1H ,T2,F10.4,T19,F10.4,T32,F10.4,T47,F10.4,T62,F10.4,
$T77,F10.4,T92,F10.4)
C
C     OPTION 5: PLOT HISTOGRAM WITH PLOTPR PACKAGE
970  IF(IC(5).LT.1) GO TO 97
C
C     COMPUTE HISTOGRAM VALUES
C
        DIV=1.0
        ICNT = 0
        IF(IC(19).EQ.1) DIV=0(1)
        DO 95 I=1,11
        ICNT = ICNT + 1
        85  Z(ICNT) = DE0(I)/DIV
        DO 94 I = 2,*K
        IF(IC(19).EQ.1) DIV=0(I)
        DO 93 J = 1,10
        ICNT = ICNT + 1
        93  Z(ICNT) = DE0(I)/DIV
        94  CONTINUE
        PM = 0.001
        IPP = 20
        K10 = 10*K
        94  IPP = IPP + 10
        IF(RM(IPP).LT.PM.AND.FN(IPP).LT.PM.AND.Z(IPP).EQ.0.1) GO TO 83
        IF(IPP.LT.K10) GO TO 84
        IP = IPP
        IF(IP.LT.K10) GO TO 92
        K11 = K10 + 1
        DO 90 I = K11,IP
        90  Z(I) = 1.0
        92  CONTINUE
        CALL INITPL
C
C     OPTION 15: NUMBER OF PAGES NEEDED FOR PLOT(DEFAULT=1 DOUBLE PG)
        NUMPAG = 1
        IF(IC(15).NE.0) NUMPAG=IC(15)
        MAXSCA = 2
        IVAR(1) = 1HY
        IVAR(2) = 5HYHIST
        IVAR(3)=5HYGAVM
        IVAR(4) = 4HYRAY
        ITP(1) = 6H0HSEPV
        ITP(2) = 5H0 4GA

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ITP(3) = 64HVA AN
ITP(4) = 6H0 RAYL
ITP(5) = 6HFLIGH P
ITP(6) = 6HLATS
ITX(1) = 6HVALUE
ITX(2) = 6HDF PAN
ITX(3) = 6HDDC VA
ITX(4) = 6HRIABLE
ITX(5) = 6H
ITX(6) = 6H
ITY(1) = 6HDENSIT
ITY(2) = 6HY
ITY(3) = 6H(AUP,/)
ITY(4) = 6HBTIN)/*N
ITY(5) = 6H*BTIN=4
ITY(6) = 6HIRTH

C
C   TC(5)=2 WRITE VALUES BEFORE PLOTTING
  IF(IC(5).EQ.2) WRITE(6+9190)(IVAR(I),I=1,4)
  9190 FORMAT(1H1,1Y,A1,10Y,A5,6X,A5,6X,A4)
C
C   OPTION 6: DISTRIBUTIONS IN MKS UNITS
  IF(IC(6).NE.1) GO TO 80
  ITX(5) = 6H MKS
  DO 81 I = 1,IP
  XNMS(I) = VN(I)*CONVER
  ZM(I) = Z(I) / CONVER
  FNMS(I) = FM(I) / CONVER
  RNMS(I) = PN(I) / CONVER
  81  WRITE(7) (XNMS(I),ZM(I),FNMS(I),RNMS(I))
  GO TO 70
  80  CONTINUE
  DO 91 I = 1,IP
  IF(IC(5).EQ.2) WRITE(6+9191) XN(I),Z(I),FN(I),RN(I)
  9191 FORMAT(1H ,4F10.4)
  91  WRITE(7) (VN(I),Z(I),FM(I),RN(I))
  70  CONTINUE
  CALL PLOTPR(3,4,IVAR)
  97  CONTINUE
C
C   OPTION 7: TIME EFFECT PLOTS
C
  IF(IC(7).NE.1) GO TO 95
C
C   TEST TO SEE IF TIME EFFECT PLOTS ARE GOOD
C
  IFG1=0
  IFG2=0
  IF(NL>0.1) GO TO 95
  DO 91929 II=2,NL
  IF(RGX(II).NE.PGX(II-1)) IFG1=1
  91909 IF(RDX(II).NE.RDX(II-1)) IFG2=1
  IF(IFG1.EQ.0.OR.IFG2.EQ.0) GO TO 95
C
  CALL INITPL
C   OPTION 10: NUMBER OF PAGES NEEDED FOR PLOT(DEFAULT=1 DOUBLE PAGE)
  NUMPAG = 1

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      IF(IC(16).NE.0) NUMPAG=IC(16)
      NPLOT=2
      FROM(1)=0.0
      C  OPTION 11: SCALE FOR 1"INDEPENDENT VARIABLE ON PLOT( DEFAULT=.1)
      SCA(1)=0.05
      IF(IC(11).NE.0) SCA(1)=10**IC(11)
      MAXSCA=2
      ITX(1)=6HTIME E
      ITX(2)=6HFFFECT
      ITX(3)=6HPLOTS
      ITX(1)=6HTIME I
      ITX(2)=6HN HOUR
      ITX(3)=6HS
      ITY(1)=6HDESIGN
      ITY(2)=6H VALUE
      ITY(3)=6H
      ITY(4)=6H
      IVAR(1)=4HTIME
      IVAR(2)=6HYSUPR
      IVAR(3)=6HYSURG
      IVAR(4)=6H
      C  OPTION 8: TIME EFFECT PLOTS IN MKS UNITS
      C
      IF(IC(8).NE.1) GO TO 60
      IVAR(2)=6HYD-MKS
      IVAR(3)=6HYG-MKS
      ITY(3)=6H(MKS)
      DO 61 J=1,NL
      WRITE(3) (ESC(J),RDXM(J),PGXM(J))
      GO TO 65
      61  CONTINUE
      DO 65 J=1,NL
      WRITE(3) (ESC(J),RDX(J),PGX(J))
      65  CONTINUE
      CALL PLOTPR(3,7,IVAR)
      7= CONTINUE
      IF(IPLOT.EQ.0) GO TO 99
      ***** END 10000 LOOP *****

      C
      C
      ***** PLOT RESULTS *****
      C
      CALL PLOT(YLL,0.,-3)
      CALL HISTO
      YN(IP+1)=FIRSTY
      XN(IP+2)=DELX
      FN(IP+1)=FIRSTY
      FN(IP+2)=DFLY
      RN(IP+1)=FIRSTY
      RN(IP+2)=DFLY
      CALL LINE(XN,FN,IP,1,0,1)
      IF(7RAY.EQ.0) GO TO 10004
      CALL LINE(YN,RN,IP,1,5,3)
      10004 CONTINUE
      GO TO 99
      10003 IF(IPLOT.EQ.0) STOP
      CALL PLOT(YLL,0.,999)
      STOP
      END

```

```

SUBROUTINE GETCMU(AT,U,F,S,YBAR,GAMMA,OCF,F,IC)
COMMON/PAParams/M,AMBDA,C,R
COMMON/OPR/ICP(20),ICK1
COMMON/IINPUT/X(100),ELL(7),N,ALP,D(100),ISTRD,IPLOT,TRAY,IC(20)
COMMON/PRNT/FRS(100),FTH(100),DEN0(100),DEN1(100)
S,FTH0(100),DENTR(100),CONVER,CAL
COMMON/HIST/FINL0,K,2IN*(100),H(100),YMAX,YLEN,YLEN,CNT(5)
DIMENSION X"KS(100),DFL"KS(100)
DIMENSION DHEAD(100)
REAL M
DATA DHEAD/100*2HD/
IF(IC(20).EQ.1) GO TO 5
C
C           ***** SOLVE FOR " USING WEGSTEIN ITERATIVE METHOD *****

C
C
C           ***** START ITERATION LOOP *****

C
3  IF(PSI1(M))
      U = PSI2(M)
      M=M-U/E+*1.5
      CALL WEG(2,M,0,NC)
      WRITE(E,461) M
461 FORMAT(1HD*2HM=,E10.4)
      NE = NC+1
      IF(NE.LT.50) GO TO 2000
      WRITE(E,200)
200  FORMAT(74H NO CONVERGENCE OBTAINED FOR YSUB)
      ICK=1
      RETURN
2000 IF(M.GE.0.) GO TO 2001
      M = .0001
2001 CONTINUE
      IF(NE,50,0) GO TO 3
C
C           ***** END ITERATION FOR M *****

C
C           ***** COMPUTE C *****

C
2  CONTINUE
      U1 = PSI(M)
      U2 = PSI1(M)
      C = SQRT(U2/S)
      IF(T.GE.0.) C=-C
      OOC = 1./C
C
C           ***** COMPUTE LAMBDA *****

C
4  AMBDA = EXP(-YBAR+U1*00C)
5  CONTINUE
      ICK1=0
      IF(IC(20).EQ.1) CALL MLHCMU
      IF(ICK1.EQ.1) RETURN
      WRITE(E,101) CNT

```

```

101  FORMAT(1H1.5A10)
      CMA = 1.-ALP
      WRITE(6,190) CONVER ,CAL
180  FORMAT(1H ,8HCONVER= ,F10.3,5X,2CHCALIBRATION FACTOR= ,F10.3)
      COMPUTE XMKS(1) AND DEL(MKS)
      DO 17 JI=1,K
      DELMKS(JI)=D(JI)*CONVER
17    XMKS(JI) = X(JI) * CONVER
      FIND MKS VALUE OF YEAR
      YBARMK = YBAR + K*ALOG(CONVER)
      WRITE(6,300) N,K,ALP,          DMA
300  FORMAT(//1X,3H#= ,I5.2X,3HK= ,I5.2X,7HALPHA= ,G10.4+2X,
      ,2X,9H1-ALPHA= ,G10.4)
      WRITE(6,297)
297  FORMAT(//1X,*DEL ARRAY*)
      WRITE(6,298) (DHEAD(I),I,D(I),I=1,K)
298  FORMAT(1H ,9(42,I2,14),F7.4+2X))
      IF( CONVER.NE.1) WRITE(6,294)
294  FORMAT(1H ,*DEL ARRAY(MKS)*)
      IF( CONVER.NE.1) WRITE(6,298) (DHEAD(I),I,DELMKS(I),I=1,K)
      WRITE(6,301) YBAR,YBARMK,S,T
301  FORMAT(1Y,6HYSR= ,G12.4+2X,11HYSP(MKS)= ,G12.4+2X,
      ,3HS= ,G12.4+2X,3HT= ,G12.4)
203  FORMAT(1Y,3HM= ,G12.4+3Y,3HLC= ,G12.4+3Y,3HCD= ,G12.4+3Y)
C
C      ***** EVALUATE GAMMA FUNCTION *****
C
C      IF(IEP.EC.125.0P.IER.EG.130) WRITE(6,500) IEP
500  FORMAT(4H ERROR IN GAMMA FUNCTION COMPUTATION.  ERROR= ,I3)
      IF( IC(15).EG.1.4ND. IC(20).NE.1) WRITE(6,250) AMRDA,M,C
250  FORMAT(1H ,2EHSTACY-M1HRAAN PAR.  AMRDA= ,F16.5,4H  M= ,F14.5,
      ,4H  C= ,F16.5)
      CALL DIST
      CMA=1.-ALP
      DO C=1./C
      CALL MGAMMA(M,GAMMA,TER)
      AMKS=AMRDA/CONVER
      PMKS = R * (CONVER**2)
      MET=10F - STACY/M
      HOD=PHIHPAV
      IF( IC(20).EG.2) MET=10HMAXIMUM LT
      IF( IC(20).EG.1) HOD=PHKELIMOD
      WRITE(6,302) " ,C=AMRDA,AMKS,R,PMKS,MET,HOD
302  FORMAT(//1X,3HM= ,G12.4+2Y,3HCD= ,G12.4+2Y,3HLC= ,G12.4+2Y,3HSP(MKS)= ,G12.4+2Y,3HCD(MKS)= ,G12.4+2Y,3HSP(MKS)= ,G12.4+2Y,3HT= ,VIA ,A10+8,7H METHOD)
      WORD=8+PAYLEIGH
      WRITE(6,110) WORD,WORD
110  FORMAT(1T3,PHVALUE OF ,T19,RHVALUE OF ,T20,PHMEASURED,T51,
      T48 ,T65,5HGMAMMA,T62,8HMFASUPED,T29,AP          ,T114,5HGMAMMA
      3.,T4,5H RANDOM,T19,5H RANDOM,T35,5H FREQUENCY,T51,5H PREDICTED,T66,
      5H PREDICTED,T82,7HDE4SITY,T29,5H PREDICTED,T114,5H PREDICTED,/,/
      T35,5H VARIABLE,T19,5H VARIABLE,T51,5H FREQUENCY,T66,5H FREQUENCY,
      T29,7H/PIN/N,T98,7H DENSITY,T114,7H DENSITY,/,T20,5H(MKS))
      WRITE(6,120) (Y(I),XMKS(I),FDR(I),FTH(I),FTH(I),DEN(I),
      ,DETR(I),DEN(I),I=1,K)
120  FORMAT(1T3,F8.4,T19,FP.4,T32,F12.4,T49,F12.4,T63,F12.4,
      ,T79,F12.4,T96,F12.4,T111,F12.4)
      RETURN
      END

```

```

SUBROUTINE HISTC
COMMON/HIST/BINLC,NOBINS,BINW(100),ORD(100),YMAX,XLEN,YLEN,CYNT(5)
COMMON/HIST2/FIRSTX,DELX,FIRSTY,DELY
DIMENSION X(200),Y(200),RAY(5)
N2=NOBINS+1
ORD(N2)=YMAX
C      COMPLETE SCALE FOR Y-AXIS
C      ORD IS THE ARRAY OF HEIGHTS FOR THE BINS
C      YLEN IS THE LENGTH IN INCHES OF THE Y-AXIS DEFINED IN
C      A DATA STATEMENT IN THE MAIN PROGRAM TO BE 10 INCHES
C      N2 IS THE NUMBER OF POINTS IN ORD AND +1 SAYS EVERY POINT IS
C      TO BE PLOTTED
C      FIRSTY THE FIRST Y VALUE ANNOTATED ON THE Y-AXIS IS RETURNED
C      IN THE N2+1 LOCATION OF ORD
C      DELY THE SCALE VALUE - AMOUNT OF DATA PER INCH- IS RETURNED
C      BY SCALE IN THE N2+2 LOCATION OF ORD
CALL SCALE(ORD,YLEN,N2+1)
FIRSTY=ORD(N2+1)
DELY=ORD(N2+2)
TITLE=10HEXCURSIONS
TITLY=7HDENSITY
ORD(N2)=0.
C      SCALES SET UP BY AXIS TO BE USED BY LINE TO PLOT
C      POINTS - X-AXIS LENGTH IS CONSTANT AND IS DEFINED IN A
C      DATA STATEMENT IN THE MAIN PROGRAM TO BE 10 INCHES
C      THE X AXIS ANNOTATION WILL BE DONE WITHOUT THE USE OF
C      AXIS SO THAT THE NUMBERS LINE UP WITH THE BINS DRAWN
KOUNT = 1
N = 1
X(N) = 0.0
N = N+1
IF (PINLC.EQ.0.0) GO TO 5
Y = N+1
X(N) = PINLO
N2 = N2+1
5 DO 10 I=1,N2
X(I) = X(I-1)+PINW(KOUNT)
KOUNT = KOUNT+1
10 CONTINUE
CALL SCALE(X,XLEN,N2+1)
FIRSTX = X(N2+1)
DELX = X(N2+2)
UNIT = 1./DELX
C      PUT PEN AT ORIGIN
CALL PLOT(0.,0.,0.)
DO 20 I=1,N2
STEP = X(I)*UNIT
C      DRAW AXIS PORTION FROM CURRENT LOCATION TO END OF NEXT BIN
CALL PLOT(STEP,0.,2)
C      DRAW TICK MARK
CALL PLOT(STEP,-.1,2)
C      ANNOTATE AXIS WITH NUMERALS
CALL NUMBER(STEP,-.21,-.21,.1,X(I),0.0,.2)
C      RETURN PEN TO AXIS LINE WITHOUT DRAWING LINE
C      CONTINUE DRAWING AXIS AND ANNOTATING IT FOR THE REMAINDER
C      OF THE VALUES
20 CONTINUE

```

```
C      TITLE THE AXIS
CALL SYMBOL(4.25,-.32,.14,TITLEX,0.0,10)
CALL AVIS(0.,0.,TITLEY,.+7,YLEN,90.,FIRSTY,DELY)
RAY(1)=10H+RAYLEIGH
RAY(2) = 10H*DIST.
RAY(3)=10H-COMPUTED
RAY(4) = 10H*VALUE
RAY(5) = 10H
XPAGE=XLEN-2.8
YPAGE=YLEN-.5
CALL SYMBOL(XPAGE,YPAGE,.14,RAY(1),0.0,10)
YPAGE=XPAGE+1.4
CALL SYMBOL(XPAGE,YPAGE,.14,RAY(2),0.0,10)
YPAGE=XLEN-2.8
YPAGE=YLEN-1.
CALL SYMBOL(XPAGE,YPAGE,.14,RAY(3),0.0,10)
YPAGE=XPAGE+1.4
CALL SYMBOL(XPAGE,YPAGE,.14,RAY(4),0.0,10)
YPAGE=YPAGE+1.4
CALL SYMBOL(XPAGE,YPAGE,.14,RAY(5),0.0,10)
YPAGE=XLEN-2.8
YPAGE=YPAGE-.5
CALL SYMBOL(XPAGE,YPAGE,.14,CNT,0.0,50)
X(1)=BINL0
Y(1)=0.
X(2)=BINL0
Y(2)=ORD(1)
J=2
DO 1 I=1,NBINS
J=J+1
XT = X(J-1)+BINW(I)
Y(J)=EXT
Y(J)=ORD(I)
J=J+1
Y(J)=EXT
Y(J)=ORD(I+1)
1  CONTINUE
X(J+1)=FIRSTY
Y(J+2)=DFLY
Y(J+1)=FIRSTY
Y(J+2)=DFLY
CALL LINE(X,Y,J+1,0,0)
RETURN
END
```

```
FUNCTION PSI(M)
C
C          FUNCTION PSI EVALUATES THE DIGAMMA FUNCTION PSI(4) FOR
C          ARGUMENT M. PSI(M) IS THE DERIVATIVE WRT M OF THE LN OF
C          THE GAMMA FUNCTION.
C
C          REAL W
C          DATA C16/.C1126/.166666657,.007936508/
C
C          T1 = 0.
C          W = "
C          IF(.4.GT.13.) GO TO 2
C                           "LF=13
C          " = 14. - W
C          W = " + "
C          D = " - 1
C          DO 1 I = 1,N
C          1   T1 = T1 + 1. / (D + I)
C                           "GT=13
C          2   RCW = 1. / W
C          RCW2 = RCW**2
C          PSI = ALOG(W) + .5*RCW + (-1. + RCW*(-C16 + RCW2*(.1*C16 - C1126+
C          . RCW2))) - T1
C          RETURN
C          END
```

```
FUNCTION PSI1(M)
C
C      FUNCTION PSI1 EVALUATES THE TRIGAMMA FUNCTION PSI (")  

C      FOR ARGUMENT M. PSI1(M) IS THE DERIVATIVE WRT M OF THE  

C      DIGAMMA FUNCTION.
C
C      REAL M
C      DATA C16,C130,C142/.166666667,.033333333+.023819524/
C
C      T1 = 0.
C      W = M
C      IF(M.GT.13.) GO TO 2
C                           M.LF.13
C      N = 14. - M
C      V = M + N
C      D = M - 1
C      DO 1 I=1,N
C      1   T1 = T1 + 1. / (D+I)**2
C                           M.GT.13
C      2   PCW = 1./W
C      PCW2 = PCW**2
C      PSI1 = PCW*(1.+PCW*(.5+PCW*(C16 +PCW2*(-C130 +PCW2*(C142
C                           -PCW2+C130)))))+T1
C      RETURN
C      END
```

```
FUNCTION PSI2(M)
C
C
C          FUNCTION PSI2 EVALUATES THE TETRAGAMMA FUNCTION PSI  (M)
C          FOR ARGUMENT M. PSI2(M) IS THE 2ND DERIVATIVE WRT M OF
C          THE DIGAMMA FUNCTION.
C
C          REAL M
C          DATA C16,C56/.166666667,.833333333/
C
C          T1 = 0.
C          M = M
C          IF(M.GT.13.) GO TO 2
C          M.LE.13
C
C          N = 14.-M
C          J = M + N
C          D = M - 1
C          DO 1 I=1,N
C          1   T1 = T1 + 2./(D + I)**3
C          M.GT.13
C
C          RCW=1./M
C          RCW2 = RCW**2
C          PSI2 = RCW2*(-1. +RCW*(-.5 +RCW2*(C16 +RCW2*(-C16
C          +RCW2*(.3 -C56*RCW2)))))) -T1
C
C          RETURN
C          END
```

```

SUBROUTINE WFG(I,XNP1,J,Y)
      WFGSTEIN ITERATIVE SOLUTION METHOD FOR X = F(X)
      N+1      N

INITIALIZATION
1 IF(I,NE.1) GO TO 2
1 K = 1
2 XBN = XNP1
3 XTEND = 10.***(-J)
4 XP = XNP1
5 RETURN

FIRST ITERATION
6 IF(K,NE.1) GO TO 4
7 IF(APS(XP-XNP1)-XTEND*APS(XP)) = 5.5,8
8 XP = XNP1
9 XBNM1 = XBN
10 XBN = XNP1
11 K = 2
12 XN = XNP1
13 XNP1 = XND
14 N = 0
15 RETURN

SUCCEEDING ITERATIONS
16 XBNP1 = ((XNP1+XPNM1)-(YN+XBN))/(XNP1+XP)M1-YN-XBN)
17 IF((ABS(XP-XBNP1))-(10S(XTEND+XP))) = 5.5,6
18 XP = XNP1
19 XBNM1 = XBN
20 XBN = XNP1
21 GO TO 17
22 N = 1
23 RETURN
24 END

```

SUBROUTINE AXIS  
ENTRY LINE  
ENTRY NUMBER  
ENTRY PLOT  
ENTRY SYMBOL  
ENTRY PLOTS  
RETURN  
END

```
FUNCTION G(C)
COMMON/HIST/RINLC,K,RINW(100),O(100),YMAX,YLEN,YLEN,CNT(5)
COMMON/INPUT/X(100),ELL(7),N,ALP,D(25),ISTOPD,IPLOT,IRAY,IC(20)
OLNSUM=0.
SUM0=0.
XOSUM=0.
XLOSUM=0.
DO 10 I=1,K
SUM0= SUM0+O(I)
OLNSUM=OLNSUM+ O(I)*ALOG(X(I))
XLOSUM= XLOSUM + ALOG(X(I))*(X(I)**C)*O(I)
XOSUM= XOSUM+(X(I)**C)*O(I)
10 CONTINUE
ARG= (XLOSUM/XOSUM)-(OLNSUM/SUM0)
G= (SUM0/OLNSUM)* (ALOG(C)+ ALOG(XOSUM)+ALOG(ARG)-ALOG(SUM0)-
S PSI(1./(C*ARG)))
RETURN
END
```

```
FUNCTION F(C)
COMMON /HIST/RINL0,K,BIN(4(100),D(100),YMAX,XLEN,YLEN,CMNT/5)
COMMON /INPUT/Y(100),ELL(7),N,ALP,D(25),ISTOPP,IPLOT,IRAY,IC(20)
XOSUM=0.
XLNSUM=0.
SUM0=0.
DO 10 I=1,K
XOSUM= XOSUM+ (X(I)**C)*D(I)
XLNSUM=XLNSUM+ C(I)*ALOG(X(I))
SUM0=SUM0+C(I)
XLNSUM=XLNSUM+ ALOG(X(I)) *(X(I)**C)*D(I)
10 CONTINUE
ARG= (XLNSUM/XOSUM)-(XLNSUM/SUM0)
F= ALOG(C)+ALOG(XOSUM)+ ALOG(ARG)-ALOG(SUM0)+PSI(1./(C+ARG))
+ -(C*XLNSUM)/SUM0
RETURN
END
```

```

SUBROUTINE MLHML
C
C  MAXIMUM LIKELIHOOD ESTIMATION
C
COMMON/HIST/BINLO,K,PINW(100),^C(100),YMAX,YLEN,YLEN,CMNT(5)
COMMON/INPUT/Y(100),FLL(7),N,ALP,0(25),TSTOPP,TPLOT,IRAY,IC(27)
COMMON/IUNIT/IDB,IPF,IPS,IPCH
COMMON/PARAMS/M,AMBDAA,C,*P
COMMON/OPT/ICP(20),ICK1
REAL M
IF(ICP(1).NE.0)GO TO 10
C  PAYLEIGH PARAMETERS
C=2.
GO TO 30
10 READ(IP5,101) C
101 FORMAT(F10.0)
30 IF(ICP(2).EQ.1) WRITE(IP5,102) C
102 FORMAT(1H0,20H INITIAL GUESS FOR C,F10.5)
WRITE(IP5,555)
555 FORMAT(1H ,29H ITERATIONS FOR FINDING C ROOT)
WRITE(IP5,551)
551 FORMAT(1H ,8Y,1H C,10Y,4H(F(1))
KE=0
DO 1 I=1,57
CALL WFG(I,C,4,KE)
IF(C.LT.0.) C=ABS(C)
IF(C.EQ.0.) C=.1
ZERO=IF(C)
WRITE(IP5,211) I,C,ZERO
211 FORMAT(1H ,12,F10.5,F10.6)
IF(KE.EQ.1)GO TO 22
C=G(C)
1 CONTINUE
C  "MAXIMUM NUMBER OF ITERATIONS REACHED WITHOUT CONVERGENCE
ICK1=1
WRITE(IP5,444)
444 FORMAT(1H0,16H C ROOT NOT FOUND)
RETURN
22 CONTINUE
C  C ROOT FOUND WITHIN TOLERANCE
C  FIND M
XLOSUM=0.
XOSUM=0.
OLNSUM=0.
SUM0=0.
DO 33 I=1,K
XLOSUM=XLOSUM + ALOG(X(I))+(X(I)**C)+0(I)
XOSUM=XOSUM +(X(I)**C)+0(I)
OLNSUM=OLNSUM+ 0(I)+ALOG(X(I))
SUM0= 0(I)+SUM0
33 CONTINUE
IF(ICP(5).EQ.1) WRITE(ICP,111) XLOSUM,XOSUM,OLNSUM,SUM0
111 FORMAT(1H0,7F10.4,2X,6H XOSUM=,F10.4,2X,7H OLNSUM=,F10.4,
$2X,5H SUM0=,F10.4)
M=1./((C*((XLOSUM/XOSUM)-(OLNSUM/SUM0)))
C  FOUND M--SOLVE FOR LAMBDA
AMBDAA=((XOSUM/(M*SUM0))**(-1./C))
IF(ICP(3).EQ.1) WRITE(IP5,100) AMBDAA,M,C
100 FORMAT(1H0,42H PARAMETERS ESTIMATED BY MAXIMUM LIKELIHOOD,
$11H ESTIMATION//1H ,7H AMBDAA=,F20.6,5X+2H M=,F10.6,5X+2H C=,F10.6)
RETURN
END

```

```

SUBROUTINE DIST
REAL X
COMMON/INPUT/X(100),ELL(7),N,ALP,D(100),ISTOPD,IPLOT,IRAY,IC(20)
COMMON/PARAMS/M,AMBOA,C,R
COMMON/PNT/FORG(100),FTH(100),DEN0(100),DENT(100),FTHR(100)
t,DENTR(100),CONVER,CAL
COMMON/HIST/EINL0,K,PINW(100),H(100),YMAX,XLEN,YLEN,CNT/5)
BOTTOM = 0.0
RMIN1 = 1.0
DIV=1.0
D2=1.0
DO 10 I =1,K
C  OPTION 1A: IF ON, DENSITY = (N0./BIN)/N
C           IF OFF, DENSITY = (N0./BIN)/(N*D(I))
      IF(IC(18).EQ.0) CIV=D(I)
      IF(IC(18).EQ.1) D2=D(I)
      ARGR=(X(I)+D(I))/2.0**2./R
      R1 = EXP(-ARGR)
      ARS=(AMBOA*(X(I)+D(I)/2.0))**C
      CALL MEGAM(ARGS,M,TOP,IEP)
C  OBSERVED FREQUENCY
      FOB(S(I)=P(I)*H(I)*N
C  OBSERVED DENSITY
      DEN0(I)=H(I)*D2
C  THEORETICAL RAYLEIGH(EXPONENTIAL) FREQUENCY
      FTHR(I) = (RMIN1 - R1 ) + " "
C  RAYLEIGH(EXPONENTIAL) DENSITY
      DENTR(I) = (RMIN1 - R1)/DIV
C  THEORETICAL GAMMA FREQUENCY
      IF(IER.NE.0) GO TO 10
      FTH(I) = (TOP - BOTTOM) + "
C  GAMMA PREDICTED DENSITY
      DENT(I) = (TOP- BOTTOM)/DIV
C
      RMIN1 = R1
      BOTTOM = TOP
10  CONTINUE
      RETURN
      END

```

APPENDIX B  
SAMPLE OUTPUT

IC OPTIONS USED IN THIS RUN SET IPLOT= 0 IPAY= 0  
OPTION 5= 1  
OPTION 9= 1  
OPTION 20= 1

IPD OPTIONS USED IN THIS RUN SET  
OPTION 2= 1  
OPTION 3= 1

INITIAL GUESS FOR C= 2.00000  
ITERATIONS FOR FINDING C ROOT

C	F(C)
1 2.00000	-.004530
2 1.93817	-.004446
3 1.89815	-.000400
4 1.88827	-.000042
5 1.88712	-.000000
6 1.88712	-.000000

PARAMETERS ESTIMATED BY MAXIMUM LIKELIHOOD ESTIMATION

LAMBDA= .064866 R= .978837 C= 1.887116

```

CONVER= 1.0000 CALIBRATION FACTOR= 0.0000
P22 3F 10
N= 262 K= 14 ALPHA= 1.0000E-01 1-ALPHA= 9.9900

```

RAYLEIGH DESIGN VALUES

L=	262.0000	XRSUHD=	49.2759	XRSURD(*PKS)=	49.2759
L=	546.0000	XRSUBD=	51.0232	XRSURD(*PKS)=	51.0232
L=	1638.0000	XRSURD=	53.5310	XRSURD(*PKS)=	53.5310

L = 262.0

XSUBH4=	10.0R49	XSUBM (MK\$)=	10.0349	TXSUR0=	15.9000	TXSUBH0 (MK\$)=	15.9000
XSUBH9=	15.0186	XSUB0 (MK\$)=	15.9946				
XSUBS=	22.0135	XSUBS (MK\$)=	22.1135				
XSURG=	38.0161	XSURG (MK\$)=	38.4161	TXSURG6=	38.4000	TXSURG (MK\$)=	38.4000
XSUBGA=	38.04206	XSUBGA (MK\$)=	38.4206	TXSUBGA=	38.4000	TXSUBGA (MK\$)=	38.4000
XSUBD=	52.5000	XSUBD (MK\$)=	52.5000	TXSUBD=	52.5000	TXSUBH (MK\$)=	52.5000
XSUBDA=	52.5001	XSUBDA (MK\$)=	52.5001	TXSUBDA=	52.5000	TXSUBDA (MK\$)=	52.5000

L = 546.0

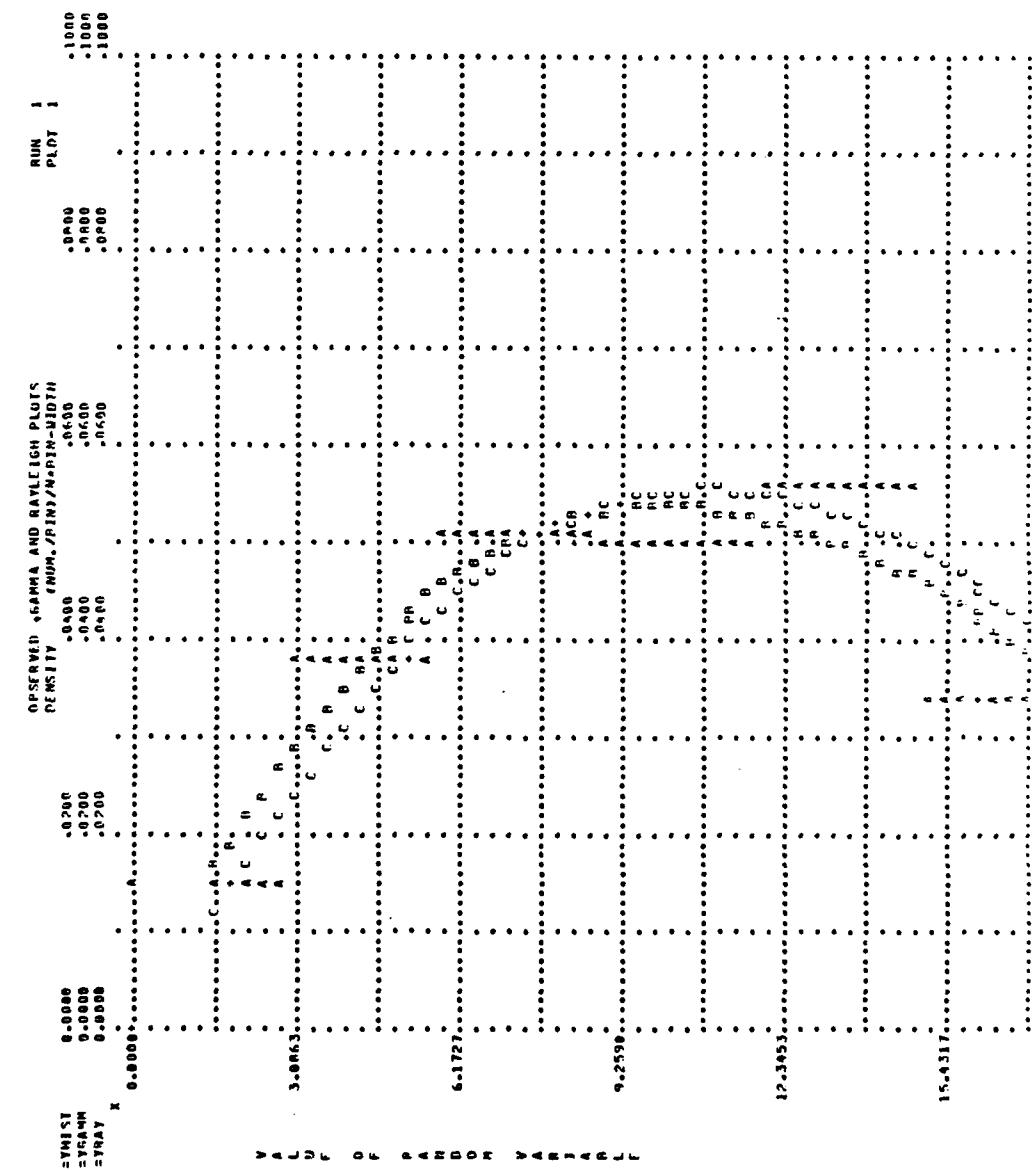
XSURM=	10.0849	XSUBM (MKS)=	10.0849	TXSURM=	15.0000	TXSUH0 (MKS)=	15.0000
XSUBU=	15.9786	XSUB0 (MKS)=	15.9906	TXSUBU=			
XSUBS=	22.1135	XSUBS (MKS)=	22.1175	TXSUBS=			
XSUBG=	40.9344	XSUBS (MKS)=	40.9344	TXSUBG=	41.1000	TXSUBG (MKS)=	41.1000
XSUBGA=	40.9670	XSUBGA (MKS)=	40.9670	TXSUBGA=	41.1000	TXSUBGA (MKS)=	41.1000
XSURD=	54.6000	XSURD (MKS)=	54.6000	TXSURD=	54.6000	TXSURD (MKS)=	54.6000
XSURDA=	54.5996	XSURDA (MKS)=	54.5996	TXSURDA=	54.6000	TXSURDA (MKS)=	54.6000

L= 1638.0

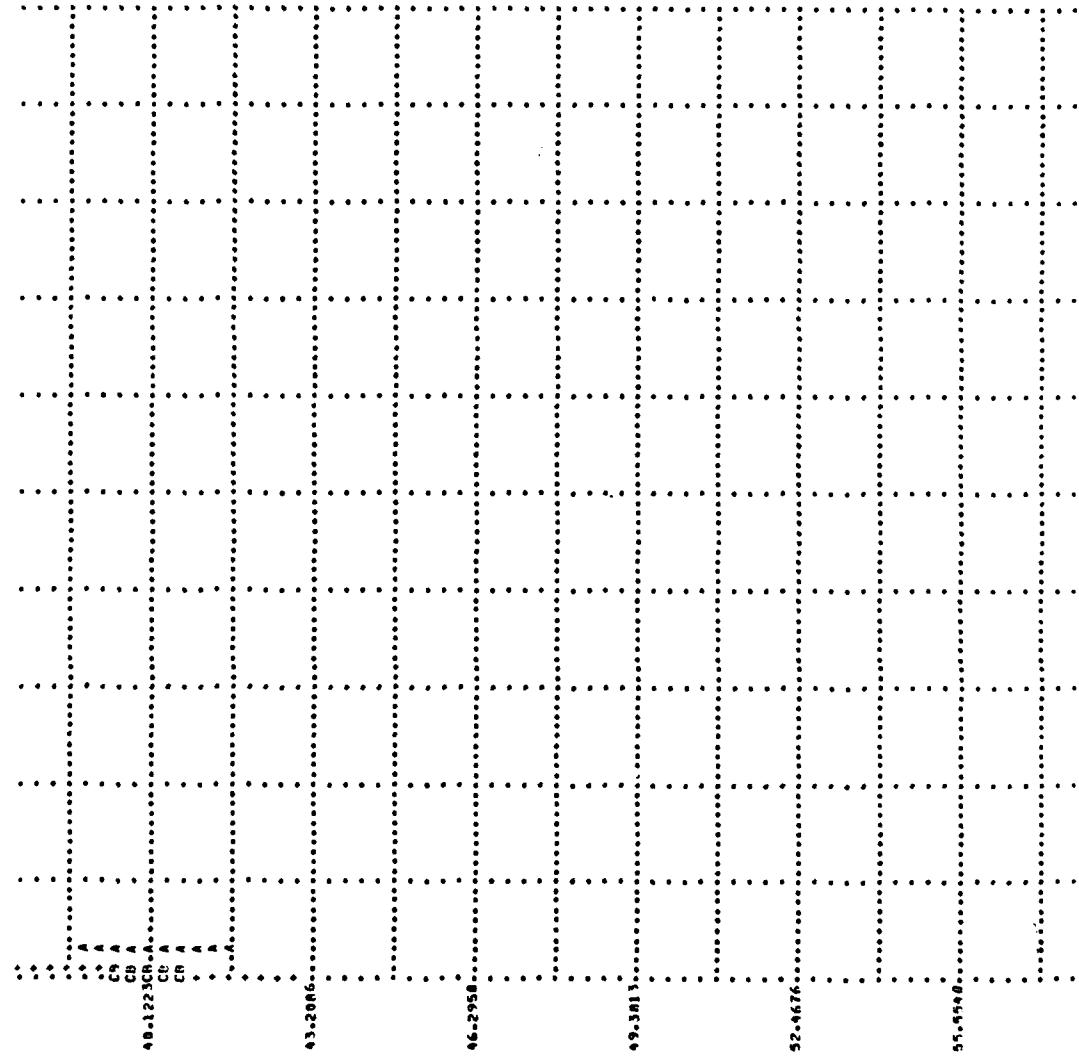
XSURH=	10.0849	XSURH (MKS)=	10.0849	TXSURH0=	15.0000	TXSURH0 (MKS)=	15.0000
XSUB0=	15.9986	XSUB0 (MKS)=	15.9986				
XSUBS=	22.1135	XSUBS (MKS)=	22.1135				
XSURG=	44.5527	XSURG (MKS)=	44.5527	TXSUBG=	44.7000	TXSUBG (MKS)=	44.7000
XSUBGA=	44.5440	XSUBGA (MKS)=	44.5440	TXSUBGA=	44.7000	TXSUBGA (MKS)=	44.7000
XSURD=	57.3000	XSURD (MKS)=	57.3000	TXSURD=	57.3000	TXSURD (MKS)=	57.3000
XSURDA=	57.3002	XSURDA (MKS)=	57.3002	TXSURDA=	57.3000	TXSURDA (MKS)=	57.3000

TIME (MIN)	TIME (HRS)	L VALUE	X SURGE	X SURGE (mks)
1.000	.0167	252.0000	52.5000	52.4161
2.0840	.0347	546.0000	54.6000	54.6000
6.2519	.1042	1678.0000	57.3000	40.9344

TIME (MIN)	TIME (HRS)	L VALUE	X SURGE	X SURGE (mks)
1.000	.0167	252.0000	52.5000	52.4161
2.0840	.0347	546.0000	54.6000	54.6000
6.2519	.1042	1678.0000	57.3000	40.9344







INITIAL GUESS FOR C = 2.00000

ITERATIONS FOR FINDING C ROOT

C F(C)

1 2.00000 -.056611

2 1.98360 -.054346

3 1.59004 -.014936

4 1.44087 -.006558

5 1.32413 -.002033

6 1.27169 -.000508

7 1.25423 -.000064

8 1.25173 -.000003

9 1.25173 -.000002

PARAMETERS ESTIMATED BY MAXIMUM LIKELIHOOD ESTIMATION

LAMBDA=

.038833  $\eta = 1.758994$   $C = 1.251726$

CONVER= P22 H 30 CALIBRATION FACTOR= 1.000

N= 311 K= 15 ALPHA= -1000E-01 1-ALPHA= .9900

DOFL ARRAY  
 D( 1)12.0000 D( 2)12.0000 D( 3)12.0000 D( 4)12.0000 D( 5)12.0000 D( 6)12.0000 D( 7)12.0000 D( 8)12.0000  
 ( 9)12.0000 D(10)12.0000 D(11)12.0000 D(12)12.0000 D(13)12.0000 D(14)12.0000 D(15)12.0000

YBAR= 3.450 YBAR(MKS)= 3.450 S= .4901 T= -.8245

M= 1.759 C= 1.252	LAMDA= .3883E-01 LAMBDA(MKS)= .3883E-01 R= 2059.	R(MKS)= 2059.
VIA MAXIMUM LIKELIHOOD METHOD VALUE OF RANDOM VARIABLE VARIABLE (MKS)		
MEASURED FREQUENCY	RAYLEIGH PREDICTED FREQUENCY	GAMMA PREDICTED FREQUENCY
RANDOM VARIABLE (MKS)	RAYLEIGH FREQUENCY	GAMMA FREQUENCY
6.0000	.2799E+02	.2809E+02
18.0000	.5086E+02	.5689E+02
30.0000	.7912E+02	.6981E+02
42.0000	.6494E+02	.6416E+02
54.0000	.5900E+02	.5593E+02
66.0000	.4903E+02	.4744E+02
78.0000	.3141E+02	.2905E+02
90.0000	.5971E+01	.656E+02
102.0000	.1866E+01	.2460E+01
114.0000	.1866E+01	.1921E+01
126.0000	.1866E+01	.2195E+00
138.0000	0.	.5251E+01
150.0000	0.	.4264E+00
162.0000	0.	.0885E-01
174.0000	0.	.1921E+00
		.8424E-01
		.3611E-01
		.30004E-03
		.1120E+01
		.3611E-03
		.8050E-07
		.9675E-05

RAYLEIGH DESIGN VALUES

XRSUBD(MKS) =	145.9384
XRSUPD(MKS) =	152.7299
XRSUPD(MKS) =	159.9K32

L = 311.0

X SURM=	24.7272	X SURM (MKS)=	24.7272	TX SUBD=	44.4000	TX SURD (MKS)=	44.4000
X SUH0=	46.9450	X SUH0 (MKS)=	46.9450				
X SUBS=	65.4169	X SUBS (MKS)=	65.4169				
X SUB6=	128.5466	X SUB6 (MKS)=	128.5466	TX SURG=	128.4000	TX SURG (MKS)=	128.4000
X SUBGA=	128.5892	X SUBGA (MKS)=	128.5892	TX SURGA=	128.4000	TX SURGA (MKS)=	128.4000
X SURD=	192.0000	X SURD (MKS)=	192.0000	TX SURD=	192.0000	TX SURD (MKS)=	192.0000
X SURDA=	192.0003	X SURDA (MKS)=	192.0003	TX SURDA=	192.0000	TX SURDA (MKS)=	192.0000

L = R33.0

XSUBM=	24.9272	XSUBM(MKS)=	24.9272	TXSUBB=	44.4000	TXSUBD(MKS)=	44.4000
XSUBD=	44.4000	XSUBD(MKS)=	44.4000	TXSUBC=	65.4169	XSUBS(MKS)=	65.4169
XSUBS=	65.4169	XSUBS(MKS)=	65.4169	TXSUBG=	143.0865	XSUBG(MKS)=	143.0865
XSUBG=	143.0865	XSUBG(MKS)=	143.0865	TXSUBG=	143.0789	XSUBGA(MKS)=	143.0789
XSUBGA=	143.0789	XSUBGA(MKS)=	143.0789	TXSUBGA=	142.8000	TXSUBG(MKS)=	142.8000
XSUBD=	205.2000	XSUBD(MKS)=	205.2000	TXSUBB=	205.2000	TXSUBD(MKS)=	205.2000
XSUBDA=	205.2000	XSUBDA(MKS)=	205.2000	TXSUBDA=	205.2000	TXSUBDA(MKS)=	205.2000

L= 2499.0

XSUBM=	219.9272	XSUBM (MKS)=	24.9272	TXSUBR0=	44.4000	TXSUR0 (MKS)=	44.4000
XSUB0=	44.4050	XSUB0 (MKS)=	44.9450				
XSUPS=	65.0169	XSUPS (MKS)=	65.4169				
XSUBG=	158.6699	XSUBG (MKS)=	158.6699	TXSUBG=	158.4000	TXSURG (MKS)=	158.4000
XSUBGA=	158.6588	XSUBGA (MKS)=	158.6588	TXSUBGA=	158.4000	TXSURGA (MKS)=	158.4000
XSUBD=	219.6000	XSUBD (MKS)=	219.6000	TXSUBD=	219.6000	TXSUPD (MKS)=	219.6000
XSUBDA=	219.5997	XSUBDA (MKS)=	219.5997	TXSUBDA=	219.6000	TXSURDA (MKS)=	219.6000

TIME (MIN)	TIME (HRS)	L VALUE	X SOUND	X SURD (MKS)	X SUNG (MKS)
1.0000	*01.67	311.0000	102.0000	128.5468	102.0000
2.6785	*04.46	813.0000	205.2000	143.0965	205.2000
8.0354	*13.39	2499.0000	219.6000	158.6690	219.6000

TIME (MIN)	TIME (HRS)	L VALUE	X SOUND	X SURD (MKS)	X SUNG (MKS)
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AD-A108 523 DAVID W TAYLOR NAVAL SHIP RESEARCH AND DEVELOPMENT CE--ETC F/6 12/1  
MODIFICATIONS TO COMPUTER PROGRAM FOR PARAMETER ESTIMATION FOR --ETC(U)  
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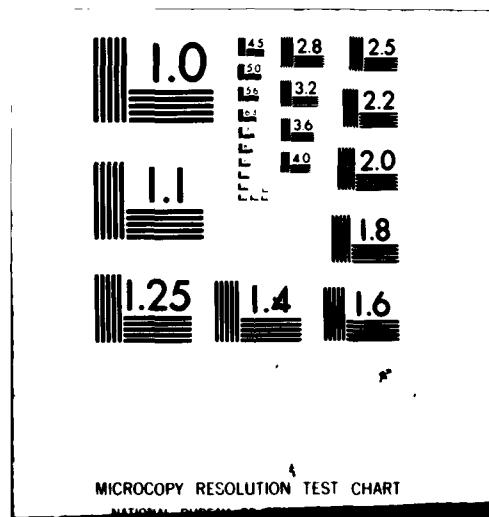
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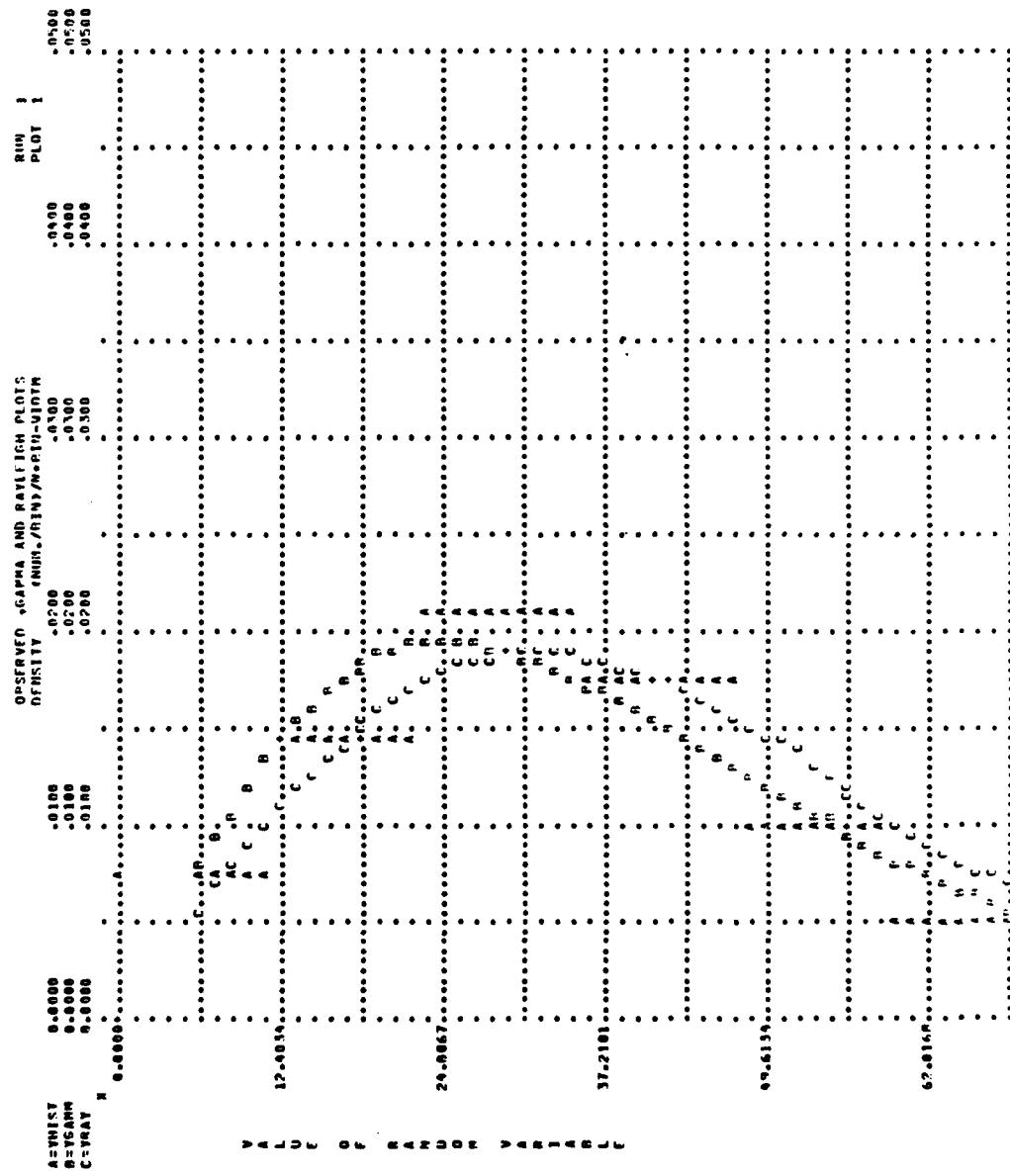
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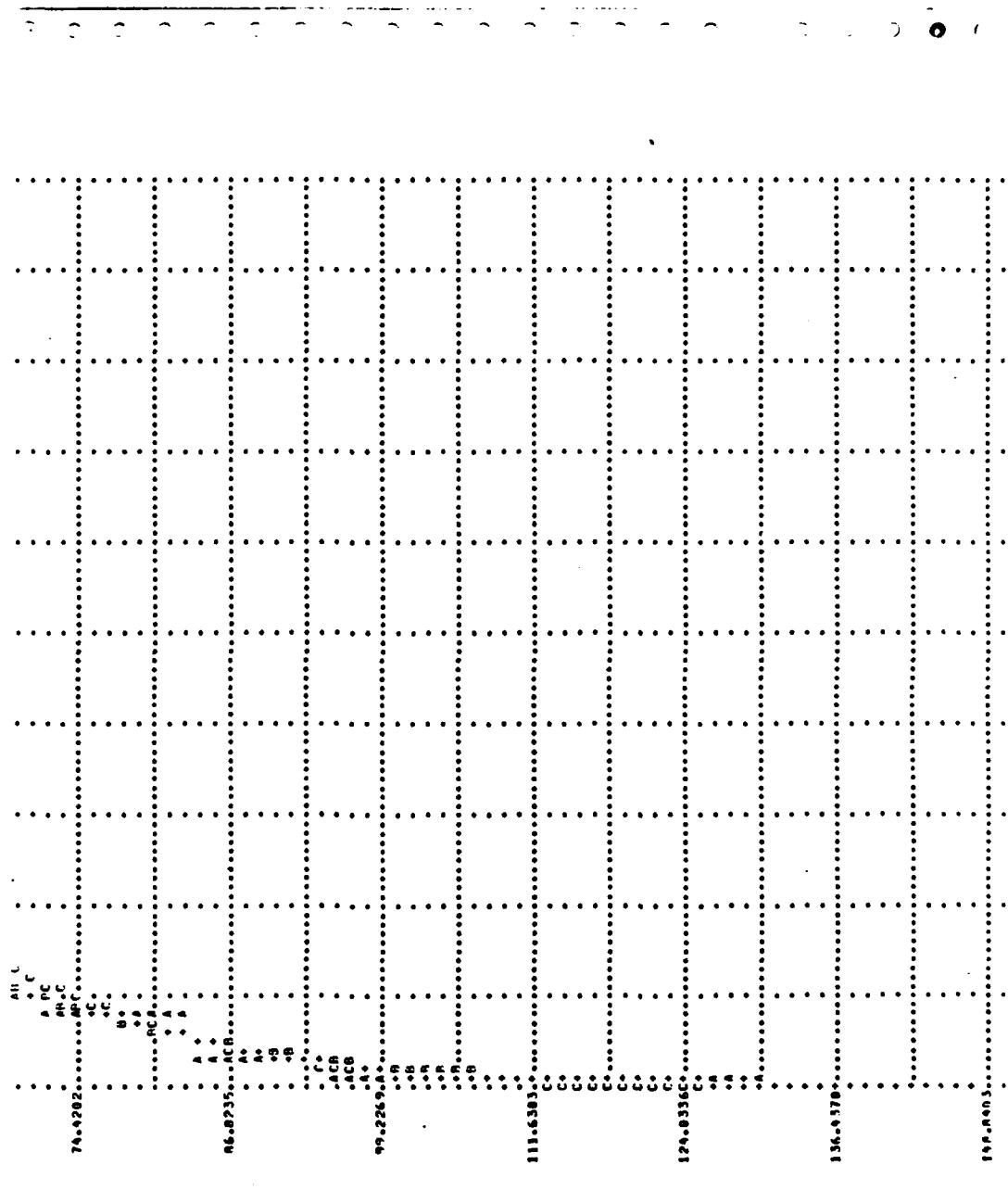
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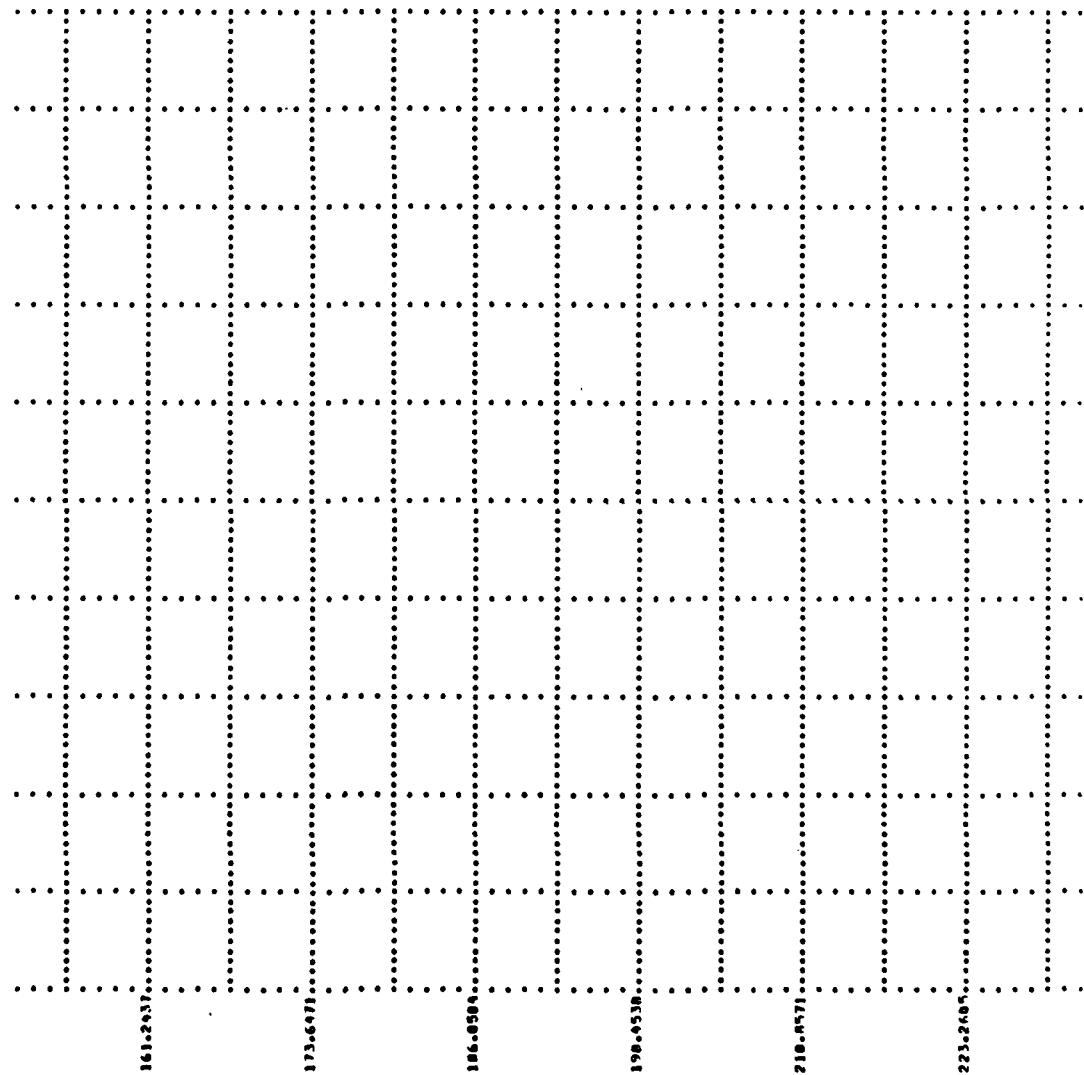


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